

ORIGINAL ARTICLE

Union membership and job satisfaction over the life course

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Abstract

We examine the relationship between union membership and job satisfaction over the life-course using data from the National Child Development Study (NCDS) tracking all those born in Great Britain in a single week in March in 1958 through to age 55 (2013). We find there is a significant negative correlation between union membership and job satisfaction that is apparent across the life-course. Lagged union membership status going back many years is negatively correlated with current job satisfaction, though its effects become statistically non-significant when conditioning on current union membership status. These results provide a different perspective to longitudinal studies showing short-term positive responses to switches in membership status. They are consistent with earlier work showing that this cohort of workers, and others before them, have persistently lower job satisfaction as union members compared to their non-union counterparts.

1 | INTRODUCTION

For many decades, the negative partial correlation between union membership and job satisfaction was taken as a given. Together with a positive union wage gap (Freeman & Medoff, 1984; Lewis, 1963) and lower quit rates among union members (Freeman & Medoff, 1984), it

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was regarded as one of the empirical regularities in labour economics. It attracted much attention subsequently because the finding appeared at odds with the objectives of trade unionism, which are to improve the terms and conditions of their members relative to what they might have been offered in the absence of a trade union.

Some studies indicated the partial correlation was biased by unobserved traits of workers or their jobs that simultaneously affected membership status and job satisfaction. Most notable were the studies indicating workers with a tendency towards job dissatisfaction sorted into union membership, perhaps because they had character traits which meant they were permanently less satisfied than non-members, and those studies suggesting union dissatisfaction arose because the unionised were in poorer jobs, or experienced worse working conditions, than their non-member counterparts, and thus sought union representation to address their problems.

Both arguments garnered support from empirical studies. Some identified poorer working conditions—such as a poor climate of industrial relations (Bender & Sloane, 1998)—as a trigger to membership. Others found the negative partial correlation between membership and job satisfaction became statistically non-significant, or even positive, when using person-fixed effects models which focused purely on the within-person variance in job satisfaction associated with switches in union membership status. These studies provided some support for the proposition that the negative partial correlation may have been driven by innate, otherwise unobserved, dissatisfaction of workers who joined unions. However, further investigation using event-study analyses revealed that, while those about to enter membership did experience substantial job dissatisfaction, this switched around on joining the union, such that they became more satisfied than they had previously been—albeit for a limited period (Powdthavee, 2011).

This literature turned another corner recently when Blanchflower et al. (2022) found the negative partial correlation between union membership and job satisfaction observed in the early studies of Freeman and Medoff and Borjas was no longer apparent in recent surveys. Indeed, the partial correlation had become positive and statistically significant. The authors speculate about potential reasons for this turnaround but one, which is apparent in the data, is that more recent entrants to union membership felt more positively about their jobs than workers born some years ago who had entered union membership decades before. The authors had no ready explanation for this switch in cohort effects underling the new-found positive correlation between union membership and job satisfaction, but they did speculate that, if the cohort trends were to continue, the correlation between membership and job satisfaction may become more positive over time.

In this study, we contribute to the literature by focusing on the relationship between union membership and job satisfaction among a cohort of workers born in Britain in the same week in 1958. We track these individuals through to age 55 to examine the association between union membership and job satisfaction over the life-course.

In doing so, we focus on the birth cohort which would have been entering the labour market around the time Freeman and Medoff (1984) and Borjas (1979) were conducting their studies. These are workers who, at least in the early years of their lives, we might expect to exhibit a negative correlation between their union membership and job satisfaction. Subsequent analyses by Blanchflower et al. (2022) for the United States using cross-sectional data suggest the link between union membership and job satisfaction remained negative for this cohort of workers. We wish to establish whether this persistent negative correlation

between union membership and job satisfaction was also apparent for a cohort of British workers born around the same time.

The rationale for our study has a second motivation: while others have used longitudinal data on individuals over time to examine links between union membership and job satisfaction, the focus has been either on comparing OLS estimates with worker fixed effects models—to isolate the effects of changes in membership on job satisfaction within individuals over time—or else to undertake event studies which estimate changes in job satisfaction shortly before and shortly after a switch in union membership status. What is lacking from the literature is an appraisal of the links between union membership and job satisfaction over the whole life-course. This is something we offer for the first time in this study. It is valuable because, by entering lagged union membership into job satisfaction equations, we are able to establish the extent to which union effects persist over the life-course. We find that union membership continues to be negatively correlated with job satisfaction over two decades later. For instance, membership at age 33 is negatively correlated with job satisfaction at age 55, and membership at age 23 is negatively correlated with job satisfaction at age 50. However, these correlations lose statistical significance when controlling for contemporaneous membership.

The rest of the study proceeds as follows. Section 2 discusses the previous literature using panel data to investigate the links between unionisation and job satisfaction, and comments on the value of a life-course investigation of the issue. Section 3 presents the National Child Development Study (NCDS) data and the methodology used to capture union effects on job satisfaction. Section 4 presents our results before we conclude and comment on implications in Section 5.

2 | LITERATURE

The earliest studies on the relationship between unionisation and job satisfaction used annual panel data on individuals from household surveys to unpick the relationship between changes in union status and changes in job satisfaction. There are a couple of strands to this study. One strand considered the potential for omitted work-related variables to explain the job dissatisfaction of unionised workers. A second considers the types of individuals who become unionised and whether they differ in unobserved ways from those who are not unionised. We take each strand in turn.

Recognising that unions are more likely to be able to organise poorer jobs in poorer working environments, and that unions seek to improve the wages and working conditions of unionised workers, analysts have investigated whether unions do indeed organise workers in poorer jobs and, having done so, whether they improve their working conditions. The first to do this were Duncan and Stafford (1980). Using the Panel Study of Income Dynamics (PSID) they examined changes in wages and working conditions for those joining unions, those leaving unions, those remaining unionised and those who remained non-unionised. They found those joining a union experienced a substantial wage hike compared to other workers, but a deterioration in their working conditions, results which they thought were consistent with unions procuring a compensating wage differential for union workers' exposure to poorer conditions. Although they don't consider employee perceptions of their jobs, the implication is that, if unionised workers are fully compensated for their poorer jobs through higher wages, their utility from the job will be similar to that obtained from a 'better' but less well-paid job. Consequently, their job satisfaction—if treated as an indicator of the utility workers derive from their jobs—would be unlikely to differ from those of similar non-unionised workers.

Their findings were immediately challenged by Leigh (1981). Using panel data from the National Longitudinal Study (NLS) he investigated changes in employee perceptions of their jobs, including 'progress at work' which, he argued (p. 1051), was akin to a job satisfaction metric, as well as pace of work, fatigue and job pressures. In contrast to Duncan and Stafford (1980), he found little evidence that these changed with union status, concluding that Duncan and Stafford's findings regarding unions procuring compensating wage differentials did not generalise to the young men in the NLS. Leigh went further, suggesting that perceived working conditions in the union sector were 'at least equal to' (p. 1054) those in the non-union sector 'as would be expected if, despite their higher wages, organised workers have a direct impact through their unions in altering disagreeable working conditions'.

Leigh underscores this point by noting a reduction in perceived job pressures among union stayers, compared to non-union stayers, while holding wages constant, consistent with union workers '[making] progress in their work' (p. 1054). His conclusion is that 'channels of collective voice encourage the expression of dissatisfaction while simultaneously restraining dissatisfied workers from leaving their employers', giving credence to the causal impact of unions in being able to extend worker tenure through the expression of, and actions to address, worker dissatisfaction.

A slightly earlier paper by Borjas (1979) is also directly pertinent to this debate. It is largely based on cross-sectional data from the National Longitudinal Survey of Mature Men aged between 50 and 64 years of age when the survey was conducted in 1971. It considered union effects on job satisfaction directly, recognising the possibility that the dissatisfaction among union members could conceivably arise due to their organising of poorer jobs. Borjas addresses this reverse causality issue in two ways.

First, Borjas runs a simultaneous equation in which union membership and job satisfaction are jointly determined, instrumenting for union membership with industry dummies. Union membership is found to be associated with higher job dissatisfaction, even when instrumented, a finding that is consistent with the proposition that unions do raise job dissatisfaction and that they do so through a voice mechanism which politicises workers as they seek better terms and conditions. Exploring this further Borjas considers union effects on job satisfaction for long- and short-tenure workers, arguing that voice effects imply greater job dissatisfaction among longer-tenure workers. This is precisely what he finds. He goes on to find that this effect is not compensated for by seniority pay. On the contrary, it appears the union wage effect diminishes with tenure. In the one panel element in Borjas' study, he examines quit rates as a function of unionisation over the period 1969–1971. He confirms the negative effect of unionisation on quits but notes that it is much stronger at lower tenure levels, becoming weaker with seniority and even positive after 25 years of tenure (p. 35), consistent with unions flattening earnings profiles as workers remain in the same job. Taken together Borjas's (1979) analyses are in line with those of Leigh (1981) in supporting the exit-voice hypothesis and run counter to Duncan and Stafford's (1980) finding that unions procure compensating wage differentials.

In his analyses for the United States using the National Longitudinal Survey of Youth (NLSY) 1979–2004, Artz (2010) emphasises the importance of past union experience for current job satisfaction, in addition to current membership status. Failure to account for past union experience effectively means that all current non-members are treated the same—as if they had never experienced unionisation. Accounting for past union experience proves important empirically since the negative association between unionisation and job satisfaction increases with union experience. Conversely, as time since last unionisation grows, so job satisfaction rebounds, becoming more positive. Artz interprets the latter as a fading of the politicisation

which, as Borjas (1979) notes, comes with union voice and bargaining. In the analyses, we present later we, therefore, incorporate lagged union status to see what effect, if any, it has on job satisfaction.

There is long-standing parallel literature for Britain using panel data.¹ Some of this has focused on union coverage—that is, whether an individual is in a job whose terms and conditions are set via union bargaining—while other papers focus on union membership. The distinction is a potentially meaningful one for Britain because, unlike the United States where there is a strong positive correlation between membership and coverage, the British industrial relations system means many of those whose pay is set via union bargaining (covered workers) may not be union members—that is, they ‘free-ride’ on union bargaining. Conversely, there are some members whose terms and conditions are not subject to collective bargaining.²

Bryson and White (2016a) examine associations between changes in employees’ satisfaction with various aspects of their job and changes in union coverage. Using panel data from the British Household Panel Survey (BHPS) (1991–2007) which follows up on individuals annually they show that when employees move into a union-covered job they receive a wage markup and, perhaps unsurprisingly, report higher pay satisfaction. Conversely, those moving from a union-covered job on average lose any markup and have significantly reduced pay satisfaction. Similar findings emerge for working hours which, they show, are also often the object of union bargaining: workers prefer shorter hours and achieve them with a switch into union coverage, resulting in higher satisfaction with working hours. However, switching into union coverage lowers satisfaction with job security, even though coverage has no effect on the perceived risk of unemployment. This is because covered employees suffer greater costs of re-employment at any given level of employment risk, partly due to the loss of the union wage markup. These findings are consistent with unions’ ability to improve employees’ working conditions and wages, resulting in improvements in those facets of the job where unions bargain successfully.

A second strand of the literature has examined the possibility that those workers who unionise are different from those who do not in ways that affect their propensity to report job dissatisfaction. If the difference is unobserved but fixed over time, then this differential propensity for dissatisfaction among unionised workers will disappear with the introduction of person-fixed effects into the job satisfaction equation. Of course, as in the case of the panel analyses seeking to account for differences in working conditions, the identification of union

¹There is also literature for Britain which relies on cross-sectional data to investigate whether union effects on job satisfaction persist when one accounts for unobserved worker or workplace effects. Bender and Sloane (1998) find the association between unionisation and job satisfaction becomes statistically non-significant when they instrument for union status with perceptions of the climate of industrial relations at the workplace. Bryson et al. (2004) obtain similar results using linked employer-employee data instrumenting for union membership using employer perceptions of the climate of industrial relations. The distinction between union coverage and union membership is found to be important in Bryson et al. (2010) who find that, having instrumented for coverage using workplace age and the workplace belonging to a multi-site firm, the link between membership and job dissatisfaction is confined to uncovered union members.

²In both countries, workers are protected from penalties associated with being, or not being, a union member. However, in the United States, in all but the ‘right-to-work states’, employees can be required to pay the equivalent of union dues if they are employed in a workplace where a union is recognised for bargaining purposes. In the United Kingdom, the closed shop (requiring an individual to be a union member to perform a particular job) was only outlawed in the late 1980s (whereas it had been illegal in the United States since 1947), such that union membership has only been an unfettered choice for all workers since the 1990s. This is worth recalling when we come on to the empirical analysis presented later because employees in NCDS may have been subject to the closed shop (depending on the industry and occupation they worked in) up until the time shortly before the NCDS fourth survey sweep in 1991.

effects is reliant on those switching union status. A series of panel studies have regressed union status on job satisfaction, treating it as exogenous. They tend to find a negative association between job satisfaction and unionisation which falls and becomes statistically non-significant, or even turns positive, when incorporating person-fixed effects which account for otherwise fixed unobserved differences across individuals.³

This is the case, for example, in the study by Bryson and White (2016b) using 17 sweeps of BHPS data for the period through to 2008: negative associations between union coverage and satisfaction with pay and hours become positive and statistically significant with the incorporation of person fixed effects.⁴ Blanchflower et al. (2022) find similar effects when comparing OLS and person fixed effects models for unionisation effects on overall job satisfaction over the period 1996–2014 using panel data combining BHPS with its successor Understanding Society (USoc).⁵ The implication is that, at least among employees in the 1990s and 2000s in Britain, un-ionised workers had a higher propensity for dissatisfaction than their non-union counterparts, a propensity which may have influenced their union status in the first place.

Conscious of the potential endogeneity of switches in union coverage status, Bryson and White (2016b) distinguish between different paths into coverage. They find union coverage effects on pay satisfaction are strongest where coverage is extended to an employee who remains in the same job, with the same employer. This, they argue, is likely a more exogenous switch than those accompanying a job switch, thus providing further weight to the proposition that unions may have a causal impact in raising employee job satisfaction through pay bargaining.

Another important finding in the literature is that union-induced changes in job satisfaction vary over time. This is what Artz (2010) found for the United States. It is perhaps best illustrated for Britain in the paper by Powdthavee (2011) for private sector employees in waves 5–15 of the BHPS. Using panel fixed effects estimates he shows job satisfaction declines in the year before becoming un-ionised, relative to those who remain non-union—as one might expect if dissatisfaction with one's job triggered the desire to become a union member. This is followed by an initial positive impact of being newly un-ionised on several aspects of job satisfaction. However, this positive effect only persists in the case of pay satisfaction, which is perhaps the issue which unions devote most time and effort to in bargaining with employers.⁶

This review suggests that unionisation tends to be negatively correlated with job satisfaction when treated as exogenous; the effects become non-significant and, in some cases positive when instrumenting for unionisation and when fixed unobserved differences across workers are accounted for; and that unionisation effects are most likely positive in the case of pay satisfaction, reflecting the wage bargaining efforts of trade unions which raise union workers'

³Not all effects move in the same direction. For example, in Heywood et al. (2002) study of the BHPS for the period 1991–1994 a negative union membership coefficient in a pay satisfaction regression becomes statistically nonsignificant with the incorporation of person fixed effects. However, negative associations between membership and satisfaction with 'relations with the boss' and 'work itself' persist with the incorporation of person-fixed effects.

⁴In their study of BHPS data for 1995–2008 Green and Heywood (2014) find the introduction of worker-fixed effects systematically reduces the size of the negative regression coefficients for unionisation with respect to satisfaction with all aspects of the job.

⁵The results are robust to the use of union membership or a dummy variable where unionisation is either membership or union coverage.

⁶Bryson and White (2016b) also test for the persistence of union positive effects on job satisfaction and find that the effects of union coverage on pay satisfaction do not persist over time.

wages above those set in the market. However, there is some uncertainty as to how long these union effects persist, and there may also be a causal impact of unionisation inducing job dissatisfaction through the politicising effects of union voice and bargaining.

The bulk of this literature focuses on short-run effects of unionisation—usually change over 1 year or, in some cases, a few years in succession. There is no literature investigating union effects over the life-course, though there are suggestions in the work by Borjas (1979) that union wage effects diminish with tenure, leading to greater job dissatisfaction among older unionised workers. Furthermore, only recently in the work of Blanchflower et al. (2022) has it become apparent that union associations with job satisfaction differ markedly by birth cohort, as well as age. Among those cohorts born more recently unionisation is more positively associated with job satisfaction: unionised employees born in the 1950s and 1960s have lower job satisfaction than their non-union counterparts, a difference that persists over time.

It is possible that the association between union membership and job dissatisfaction observed in the literature is driven by a third, usually unobserved, factor namely one's family origins which potentially generate both a greater propensity for union membership and job dissatisfaction. We know from other research that union membership is passed inter-generationally (Bryson & Davies, 2019) but little is known about one's social origins at birth and the propensity for job satisfaction. We are able to test this proposition with these data because we observe the social class of the cohort member's parents.

While the bulk of the literature on unionisation and worker wellbeing has focused on job satisfaction, recently the literature has expanded to incorporate other well-being indicators. For instance, Blanchflower et al. (2022) reported that the partial correlation between unionisation and job satisfaction was positive in post-Great Recession data, but so too was its partial correlation with other well-being measures including life satisfaction, happiness and lowered stress and worry. They also found that unions entered positively in pain and anger equations in the Gallup US Daily Tracker from 2008 to 2017.

In the remainder of this study, we turn to union effects on job satisfaction across the life-course for a British birth cohort born in 1958 who were entering the labour market in the late 1970s and who we track through to age 55 in 2013. In doing so, we contribute to the literature by establishing the association between unionisation and job satisfaction, not only contemporaneously, but with lagged membership going back many years. In doing so, we are able to establish the persistence in the partial correlation between membership and job satisfaction.

3 | DATA AND ESTIMATION

Our data are the British NCDS, a nationally representative cohort study of all the people born in England, Scotland and Wales in a single week in March, in 1958 (Pearson, 2016). In addition, 798 immigrants born in that week but outside Great Britain and 345 who were born in Great Britain, but who were not in the initial perinatal mortality study (PMS) were later added. In total there are 18,558 study participants of whom 18,059 are white. There were 395 stillbirths and 278 neonatal deaths in the first 28 days of life and by age 55, in NCDS9, there had been an additional 987 deaths. In addition, there was an attempt made in 2013 to contact those once in Britain but now living abroad, to invite participation in a web-based follow-up at age 55 (Dodgeon et al., 2014). At NCDS9 1286 study participants had emigrated.

We track the job satisfaction and union status of those in employment at ages 23, 33, 42, 46, 50 and 55—respectively sweeps 4 through 9 of the data. All but the telephone survey at age 46

involved a home interview. The cohort started as the PMS with 17,415 births covering 98% of births in that week. Individuals frequently disappear in one sweep and reappear in a subsequent one. In NCDS4 at age 23 there were 12,537 observations; 11,469 in NCDS5 at age 33, 11,419 in NCDS6 at age 42; 9534 in NCDS7 at age 46; 9790 in NCDS8 at age 50 and 9137 in NCDS9 at age 55. At age 55 approximately 86% of the sample had responded to NCDS4.

We present models estimating three sets of dependent variables. First, we estimate job satisfaction equations at sweeps 4 (age 23), 7 (age 46), 8 (age 50) and 9 (age 55). In most of the analyses, the dependent variable is overall job satisfaction, but we also present satisfaction with facets of the job at sweep 4 (age 23) which is the only age at which this richer set of job satisfaction variables is available. We everywhere restrict the sample in each sweep, including lagged values, to workers only. Although the satisfaction variables are ordinal, running from very dissatisfied to very satisfied, we run OLS estimates.⁷ The models are run for single sweeps of data and incorporate the same set of controls throughout, namely gender, region and education, together with employment status (part-time employee, self-employed). We also present extended models which incorporate the cohort member's social class, as indicated by the occupation of his or her father at birth.

Union status is captured by the cohort respondent's union membership status. We experiment with contemporaneous union membership, together with lagged union membership in previous waves to assess the persistence of union membership effects. Throughout estimates are unweighted.

Second, we run log hourly wage equations for employees at sweep 4 (age 23) and sweep 8 (age 50) to establish the size of the wage gap between union members and non-members.

Third, when employees are aged 50 in sweep 8 they are asked a broader range of subjective wellbeing questions, such as how much pain they feel at work, and whether they would pack in their job if they disliked it. These models control for social class, region and education, as well as the cohort member's gender and current employment status (part-time, self-employed). These models capture the partial correlation between five well-being metrics and contemporaneous union membership status.

By age 23, of the 12,536 individuals who responded to the NCDS4 survey 73.6% were in employment. Of these, 44.3% were union members. Recall, this was 1981 when union membership in Britain was close to its peak. Union membership was 13.2 million in 1979; 12.6 million in 1980 and 12.3 million in 1981. It fell dramatically afterwards as examined by Disney (1990) and Disney et al. (1995) such that, by 2013—when we last see our birth cohort at age 55—membership had dropped to 7.0 million.⁸ This is reflected in the NCDS birth cohort: union membership had halved to 21.3% by 2008 when they were aged 50.

In addition, unemployment was rising quickly under Margaret Thatcher's premiership. The United Kingdom's monthly unemployment rate rose steadily during 1981 from 8.6% in January through to 10.3% in December, a rate it has never reached since.⁹

⁷Throughout results are robust to the use of an ordered probit estimator.

⁸It continued to fall until 2016 when it stood at 6.23 million, rising to 6.4 million in 2019 ('Trade Union Membership, UK, 1995-2019: Statistical Bulletin', 27 May 2020, Department for Business, Energy and Industrial Strategy).

⁹Being part of a large cohort entering the world of work can reduce future prospects (Korenman and Newmark, 2000). We looked at the UK mid-year population estimates and found that the number age 23 in 1981 was 1.48% of the population, which compares with an average from 1971 to 2014 of 1.42%. The size of the age 23 cohort was on a rising trend from 1.36% in 1978 to a peak of 1.88% in 1988 before falling back to 1.15% in 2000. This compares with 1.2% in 2020. The number age 23 in 1981 was greater than the number in 2020 (833,717 vs. 815,851) despite a rise in the population of by nearly 11 million from 56.3 million to 67.2 million. Source: UK Population Estimates, 1851-2015, ONS July 2015 and Census Bureau International Database.

4 | RESULTS

Table 1 recovers the partial correlation between union membership at age 23 and five measures of job satisfaction using OLS estimates. The first is an overall measure of satisfaction with one's current job, with a 5-point response scale ranging from 'very satisfied' to 'very dissatisfied'. The four other job satisfaction variables relate to facets of the job, namely pay, physical conditions, job organisation and interest in the job. Each equation includes controls for gender, part-time and self-employed status, region and education levels (There is no age control since all are aged 23).

At age 23, 30% of employees reported being 'very satisfied' with their current job, and another half (49.8%) reported being 'satisfied'. Only 3% were 'very dissatisfied'. Nearly four-fifths of employees continued to report being either 'satisfied' or 'very dissatisfied' with their jobs in the other three sweeps when job satisfaction is asked, namely at ages 46, 50 and 55.

We had anticipated that the partial correlation between union membership and overall job satisfaction would have been negative and statistically significant given the earlier literature identifying such a relationship for this birth cohort, albeit in the United States. In fact, although the coefficient in column 1 of Table 1 is negative it is far from statistically significant (with a *t*-statistic of 0.27). However, a clear pattern emerges in columns 2–5 when we consider the association between union membership and satisfaction with various facets of one's job. Union members rate their jobs more poorly than their non-union counterparts on all non-pecuniary components (physical conditions, job organisation and job interest) but express greater satisfaction with their pay. These results are very much in keeping with the compensating wage differentials idea initially discussed by Duncan and Stafford (1980). The final column lends further support to this proposition: the union coefficient in the overall job satisfaction equation increases 10-fold relative to column 1 and becomes statistically significant (with a *t*-statistic of 2.37) when conditioning on log hourly pay. This is reminiscent of Borjas' (1979: 31) analyses for the United States in the early 1970s when he compared union effects on job satisfaction with and without a wage control.

To establish the precise size of this union wage premium we ran some estimates of log hourly wages in for workers aged 23 (available on request). We find a union wage premium of around 20% which hardly changes with the introduction of controls for education, part-time status and region dummies.¹⁰ This estimate is somewhat larger than found in 1980 in Blanchflower (1984) for all workers.

At age 50 the raw union wage premium is 9% but it falls to zero with controls. Again, this is reminiscent of Borjas' (1979) analysis for the United States which showed union effects on wages turned negative among men in their 50s who had been in their jobs for some time.¹¹

Ideally, we would have assessed the partial correlation between union membership and job satisfaction in all NCDS sweeps but job satisfaction is not available at age 33 or age 42. It is available when cohort members were aged 46. It is notable that their job satisfaction in general rises relative to the job satisfaction of those in employment at age 23: almost half (46%) report that they are 'very satisfied' with their job. Unfortunately, union membership is not recorded at age 46. Instead, in Table 2 we consider the partial correlation between job satisfaction and past

¹⁰The union wage differential in percentage terms is obtained by taking anti-logs and deducting 1.

¹¹Although among Borjas' survey of men in their 50 and 60s back in 1971 the average union effect on wages having conditioning on tenure and other controls was nevertheless positive and statistically significant (0.0485, *t*-stat = 2.02 from Table 5).

TABLE 1 Satisfaction at age 23, workers only

	Overall	Pay.	Physical conditions	Job organisation	Job interest	Overall
Union	-0.0058 (0.27)	+0.1425 (6.42)	-0.1286 (6.12)	-0.2422 (10.06)	-0.0677 (3.27)	-0.0558 (2.37)
Male	-0.1400 (6.23)	-0.2179 (9.49)	-0.1292 (5.95)	-0.2864 (11.49)	-0.0072 (0.33)	-0.1940 (7.95)
Part-time	0.1285 (2.75)	0.2405 (5.00)	0.1240 (2.74)	0.0745 (1.43)	-0.0498 (1.11)	0.1291 (2.49)
Self-employed	0.3821 (8.27)	0.3617 (7.42)	0.1372 (3.04)	0.6873 (12.62)	0.3368 (7.63)	0.0768 (2.04)
Log hourly pay						0.2879 (8.65)
Constant	4.2322	3.4191	4.2977	3.6586	4.3309	2.3741
Adjusted R^2	0.0170	0.0285	0.0120	0.0505	0.0220	0.0215
N	9134	9087	9111	9032	9098	8162

Note: All equations include region (10) and education (15) controls. T-statistics in parentheses. Those above 1.96 denote significance at a 95% confidence interval or above.

union membership status among those at age 46, in 2004. The motivation for doing so is to assess the persistence of union effects, as explored by Artz (2010) for the United States.

As in the case of 23-year-olds, we find 46-year-olds in this cohort were more satisfied with their jobs overall working as self-employed or part-time, when compared to working as full-time employees, while men were less satisfied with their jobs than women (Table 2, column 1).

If unionised at age 42, workers were likely to be less satisfied with their jobs at age 46 than those who had been working but were not unionised at age 42 (column 2). Being unionised at age 33 was also negatively and significantly correlated with job satisfaction at age 46 (column 3). The correlation with union membership at age 23 were negative but non-significant (column 4). The findings are consistent with the proposition that duration of union experience is strongly associated with lower job satisfaction, as Artz (2010) found for the United States. That said, more recent measures of membership are salient, as indicated when we condition on both membership status at age 42 and lagged membership when aged 33 (column 5) and when we include lagged union status at ages 42, 33 and 23 (column 6).

We conduct a similar exercise in Table 3, but this time for job satisfaction when workers were aged 50. The overall job satisfaction question is identical to the one asked at age 46 and age 23. In this case, we have contemporaneous union membership: this is strongly and negatively associated with job satisfaction at age 50 (column 1). When entered alone, lagged union membership is also partially correlated with job satisfaction at age 50, whether measured at age 42 (column 2), 33 (column 3) or 23 (column 4). Again, when entered together, it is contemporaneous union membership status that remains negative and statistically significant (columns 5 and 6).

A similar pattern of results emerges in Table 4 when we focus on job satisfaction when workers are aged 55. When entered separately lagged union membership is negatively and significantly associated with job satisfaction at age 55—even union membership some 22 years earlier when aged 33. (Union membership at age 23 was excluded as it was always insignificant). However, once again, contemporaneous union membership is negative and significant and substantially reduces the coefficient on lagged membership such that union membership at age 42 becomes non-significant when it is included in column 5. Finally, in column 6 we included lagged job satisfaction at age 50: unsurprisingly it is very strongly correlated with job satisfaction at age 55 but, even here, the partial correlation between job satisfaction at age 55 and union membership at age 50 remains negative and statistically significant. The implication is that, even if one partials out dissatisfaction with a job 5 years previously, union membership continues to generate dissatisfaction with the job now.

As noted earlier, it is possible that the association between union membership and job dissatisfaction is driven by a third, usually unobserved, factor namely one's family origins which potentially generate both a greater propensity for union membership and job dissatisfaction. We can test for this in NCDS because it contains data on the occupation of the respondent's father when they were born.¹² We plug this variable into estimates of job satisfaction at ages 23, 33, 42 and 50 in Table 5. We also include an (insignificant) control for whether the respondent was not in the PMS which means father's occupation was not reported, to maintain sample sizes. We find that one's social class at birth, as indicated by one's father's occupation, does indeed significantly impact one's own job satisfaction up to 55 years later. However, the negative partial correlation with union membership remains negative and statistically significant except at age 23.

¹²The data actually refers to the occupation of the respondent's mother's husband.

TABLE 2 Overall job satisfaction at age 46, workers only

	Model 1	Model 2	Model 3	Model 4	Model 5	Model 6
Union age 42		-0.1063 (4.03)			-0.0726 (2.27)	-0.0882 (2.30)
Union age 33			-0.0688 (2.53)		-0.0259 (0.81)	-0.0316 (0.81)
Union age 23				-0.0235 (0.83)		0.0253 (0.77)
Male	-0.1668 (6.27)	-0.1733 (6.38)	-0.1683 (5.94)	-0.1681 (5.12)	-0.1816 (6.28)	-0.1828 (5.25)
Part-time	0.0174 (0.54)	-0.0030 (0.09)	0.0073 (0.21)	-0.0002 (0.01)	-0.0067 (0.19)	-0.0284 (0.67)
Self-employed	0.2849 (8.62)	0.2617 (7.58)	0.2570 (7.20)	0.2859 (7.38)	0.2564 (6.94)	0.2454 (5.80)
Constant	4.1336	4.2121	4.1626	4.1526	4.2043	4.2301
Adjusted R ²	0.0134	0.0156	0.0141	0.0152	0.0151	0.0167
N	8302	7905	7372	5651	7061	4955

Note: All equations include region (10) and education (15) controls. Those above 1.96 denote significance at a 95% confidence interval or above.

TABLE 3 Overall job satisfaction at age 50, workers only

	Model 1	Model 2	Model 3	Model 4	Model 5	Model 6
Union age 50	-0.1527 (6.10)				-0.1357 (4.75)	-0.0744 (2.02)
Union age 42		-0.0904 (3.81)			-0.0355 (1.34)	-0.0491 (1.30)
Union age 33			-0.0653 (2.67)			-0.0259 (0.73)
Union age 23				-0.0617 (2.40)		-0.0210 (0.71)
Male	-0.1448 (6.18)	-0.1493 (6.19)	-0.1264 (5.02)	-0.1729 (5.92)	-0.1478 (6.14)	-0.1479 (4.75)
Part-time	-0.0153 (0.54)	-0.0076 (0.26)	-0.0039 (0.13)	-0.0551 (1.53)	-0.0152 (0.52)	-0.0667 (1.76)
Self-employed	0.1198 (4.08)	0.1427 (4.70)	0.1287 (4.13)	0.1349 (3.91)	0.1234 (4.03)	0.0947 (2.49)
Constant	4.2468	4.2304	4.1882	4.2148	4.2529	4.2191
Adjusted R ²	0.0131	0.0110	0.0083	0.0105	0.0139	0.0117
N	8235	7736	7194	5552	7734	4830

Note: All equations include region (10) and education (15) controls. Those above 1.96 denote significance at a 95% confidence interval or above.

TABLE 4 Overall job satisfaction at age 55, workers only

	Model 1	Model 2	Model 3	Model 4	Model 5	Model 6
Union age 50		-0.1743 (5.65)			-0.1519 (4.32)	-0.1087 (3.80)
Union age 42			-0.0942 (3.33)		-0.0261 (0.81)	
Union age 33				-0.0753 (2.59)		
Male	-0.1017 (3.73)	-0.0983 (3.41)	-0.1133 (3.95)	-0.0948 (2.78)	-0.1084 (3.69)	-0.0298 (1.11)
Part-time	0.0794 (2.55)	0.0794 (2.39)	0.0732 (2.21)	0.0945 (2.78)	0.0738 (2.18)	0.0949 (3.07)
Self-employed	0.2232 (7.03)	0.1720 (5.00)	0.1915 (5.61)	0.2149 (6.16)	0.1710 (4.83)	0.1420 (4.45)
Job satisfaction age 50 (ref: very dissatisfied)						
Dissatisfied						0.2462 (2.23)
Neither						0.4367 (4.21)
Somewhat satisfied						0.7958 (8.02)
Very satisfied						1.2990 (13.13)
Constant	4.0827	4.1311	4.1057	4.0616	4.1240	3.1548
Adjusted R^2	0.0092	0.0134	0.0105	0.0112	0.0132	0.1349
N	6942	6161	6231	5868	5921	6301

Note: All equations include region (10) and education (15) controls. Those above 1.96 denote significance at a 95% confidence interval or above.

TABLE 5 Overall job satisfaction at various ages and union membership having conditioned on father's social class in 1958, workers only

	Age 55	Age 50	Age 46	Age 23
Union age 50	-0.1744 (5.77)	-0.1493 (5.81)		
Union age 42			-0.1053 (3.98)	
Union age 23				-0.0183 (0.83)
Male	-0.1017 (3.63)	-0.1368 (5.70)	-0.1726 (6.35)	-0.1293 (5.71)
Part-time	0.0770 (2.41)	-0.0167 (0.57)	-0.0043 (0.13)	0.1127 (3.49)
Self-employed	0.1778 (5.32)	0.1262 (4.15)		
Social class at birth	Yes	Yes	Yes	Yes
Clerical workers	-0.1388 (2.23)	-0.0811 (1.57)		
Personal service		0.2956 (2.94)		
Shop assistants	0.1397 (2.19)		0.1456 (1.69)	
Foremen	-0.2647 (2.87)			
Farmers	0.1535 (1.68)			0.3182 (4.03)
Farm workers		0.1295 (2.07)	0.1197 (1.70)	
Higher admin etc		0.1040 (1.97)		
Not in PMS	-0.0383 (0.64)	0.0338 (0.70)	0.0315 (0.56)	-0.0622 (1.20)
Constant	4.1529	4.2449	4.2021	3.8653
Adjusted R^2	0.0171	0.0150	0.0157	0.0126
N	6509	8235	7905	9102

Note: All equations include region (10) and education (15) controls. Variables are measured at each sweep so in column 1, part-time and self-employed are all at NCDS9. Those above 1.96 denote significance at a 95% confidence interval or above.

Following Blanchflower et al. (2022) we sought to establish the association between union membership and a broader range of wellbeing metrics (Q6–Q10) in NCDS which were collected in 2008 when cohort members were aged 50. Results are reported in Table 6. We find union members at age 50 are likely to report more pain (column 1) and more pain at work (column 2) than their non-union counterparts, reminiscent of the Gallup Daily Tracker results mentioned above. In addition, union members were more critical of management than non-members: they were more likely to agree with the statement that managers always try and get the best of workers (column 3).

Of note is that in all three equations the occupation of the respondent's father when they were born, as recorded in the 1958 PMS, has an impact on well-being 50 years later. The final column indicates that, despite expressing greater dissatisfaction with their jobs and management at age 50 compared to their nonmember counterparts, they were nevertheless more attached to paid work than non-members: they were less likely to agree that they would pack up a job they disliked, even if they had no other to go to. Father's occupation was not a significant influence in the final column so was omitted. These responses provide some insights into why union members invest in unionisation in the first place: they prefer to voice to improve their job as opposed to simply putting up with a bad situation.

TABLE 6 Well-being equations at age 50, workers only

	Pain	Pain at work	Managers versus workers	Won't pack in disliked job
Union age 50	0.0743 (2.17)	0.0641 (2.64)	-0.2355 (8.34)	0.0731 (2.72)
Male	-0.1281 (4.00)	-0.0862 (3.79)	-0.2598 (9.85)	-0.1192 (4.74)
Part-time	0.0438 (1.13)	0.0626 (2.28)	-0.0797 (2.50)	-0.2421 (7.95)
Self-employed	0.0701 (1.71)	0.0587 (2.02)	0.1440 (4.28)	-0.3408 (10.63)
Social class at birth	Yes	Yes	Yes	No
Semi-skilled	0.1124 (2.05)		-0.1666 (3.69)	
Admin, prof, manager	-0.1301 (2.53)	-0.1009 (2.77)	0.1982 (4.69)	
Shopkeepers	-0.2222 (2.22)	-0.1570 (2.22)	0.1293 (1.57)	
Shop assistants			0.1205 (2.03)	
Personal service			0.0686 (2.81)	
Farmers			0.2563 (3.00)	
Foremen	-0.2509 (2.34)	-0.1174 (1.54)		
Higher admin etc	-0.1667 (2.35)		0.2067 (3.52)	
Not in PMS	-0.0767 (1.12)	-0.0661 (1.37)	-0.0755 (1.34)	
Constant	2.6296	1.6470	2.0833	4.0105
Adjusted R^2	0.0154	0.0156	0.1157	0.0344
N	7432	7437	7419	7417

Note: Q7. How much bodily pain last 4 weeks—none (24); very mild (25); moderate (26); severe (21) very severe (32)?

Q8. How much pain interfere with work last 4 weeks—not at all (24); slightly (26); moderate (24); quite a bit (26) very severe (29)?

Q9. Management always tries to get the better of workers—strongly agree (34); agree (27); neither (23); disagree (22) disagree strongly (17)?

Q10. Would pack in disliked job even if no other job to go to—strongly agree (16); agree (18); neither (21); disagree (26) disagree strongly (26)?

Union in NCDS8% in parentheses.

All equations include region (10) and education controls (8). Social class at birth refers to mother's husband. Excluded category is skilled workers. Those above 1.96 denote significance at a 95% confidence interval or above.

5 | CONCLUSIONS

What is lacking from the literature is an appraisal of the links between union membership and job satisfaction over the whole life-course. This is something we offer for the first time in this study using data from a cohort of workers born in a single week in March 1958. We examine the relationship between union membership and job satisfaction over the life-course through to age 55 (2013). This is valuable because, by entering lagged union membership into job satisfaction equations, we are able to establish the extent to which union effects persist over the life-course.

We find that union membership continues to be negatively correlated with job satisfaction over two decades later. For instance, membership at age 33 is negatively correlated with job

satisfaction at age 55, and membership at age 23 is negatively correlated with job satisfaction at age 50. Furthermore, these associations are robust to the inclusion of potential confounding variables, such as parental social class, which are often absent from other studies. However, these correlations between past membership and current job satisfaction lose statistical significance when controlling for contemporaneous membership.

These results provide a different perspective to longitudinal studies showing short-term positive responses to switches in membership status. They are consistent with Blanchflower et al. (2022) earlier cross-sectional analyses for the United States showing that this cohort of workers, and others before them, have persistently lower job satisfaction as union members compared to their non-union counterparts. This contrasts with later cohorts where unionised workers actually have higher job satisfaction than their non-unionised counterparts.

Why is it that, among this cohort, union membership is associated with lower overall job satisfaction across the life-course? One possibility is that the sorts of jobs organised by trades unions at that time were drawn from the lower end of the job quality distribution, either because they were concentrated in heavy industry and manufacturing, or in 'dirty' or 'dangerous' occupations. We get a hint of that from our analyses of satisfaction with different facets of the job among those aged 23: union members express greater dissatisfaction than non-members about all non-pecuniary aspects of the job but are more satisfied about their pay than non-members. This is because, as we show, unions are successful in procuring a compensating wage differential for the poor job quality their workers face. Duncan and Stafford (1980) argued the situation was similar among workers in the United States in the 1970s. That differential appears to have fully compensated union members for their poor-quality jobs at age 23, but not subsequently.

A second possibility is that the higher dissatisfaction of unionised workers reflects the operation of trade unions as an effective voice for their members, in accordance with the exit-voice model described in Freeman and Medoff (1984). The prospect of voicing out their discontent means unionised dissatisfied workers are more likely to stay in their job compared to their non-unionised counterparts, which is precisely what we find among those aged 50 (Table 6, final column).

But why has this association persisted, when later cohorts of union members express greater satisfaction with their work than non-members? It may be that, despite the changing composition of the unionised workforce associated with the decline in heavy industry and its replacement by services and public sector work, this early cohort may have become 'stuck' in poor quality jobs, incapable of moving onwards and upwards to better parts of the labour market? This is a distinct possibility, particularly in Britain because, as we noted earlier, this cohort were facing a huge economic recession in their early careers. We know that entering the world of work in a recession is particularly hard, potentially limiting future career progression (Kahn, 2010).

A second possibility is that this cohort's experience of unionisation is fundamentally different from later cohorts, and that this experience was fashioned early on in their careers. Recall, union density was at its peak in Britain when these workers entered the labour market. And unions were militant. Union recourse to strike action was common and, often, painful for union workers. The working days lost due to industrial action in the 12 months to April 1980 was 32.15 million—the highest for any 12-month period since records began in 1931. It peaked again shortly afterwards, in 1984/85 with the Miners' Strike. It is conceivable that these experiences accentuated the politicisation effect of unionisation, something which Borjas (1979) and Artz (2010) identified as important in contributing to a union job dissatisfaction

differential, which remained with this cohort of workers through the remainder of their working lives. This might also explain why it is that these workers continue to express dissatisfaction with their jobs, even years later, after their union membership, as indicated by the significance of lagged unionisation in our estimates.

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CONFLICT OF INTEREST

The authors declare no conflict of interest.

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