

## Supplementary data

**Table S1 Odds ratio for association between attrition and age, sex and comorbidity count by index condition**

| Index condition              | Comorbidity count <sup>1</sup> | Age <sup>2</sup> | Sex <sup>2</sup> |
|------------------------------|--------------------------------|------------------|------------------|
| Atrial Fibrillation          | 1.12 (1.02-1.24)               | 1.14 (0.95-1.51) | No sex model     |
| Axial Spondyloarthritis      | 1.11 (0.99-1.27)               | 1.06 (0.83-1.40) | 0.97 (0.70-1.26) |
| Benign Prostatic Hyperplasia | 1.13 (1.04-1.27)               | 1.13 (0.95-1.48) | No sex model     |
| Chronic idiopathic urticaria | 1.08 (0.95-1.18)               | 0.97 (0.73-1.15) | 0.99 (0.77-1.25) |
| Dementia (any)               | 1.10 (1.01-1.19)               | 1.08 (0.90-1.32) | 0.89 (0.71-1.06) |
| Diabetes Mellitus, Type 2    | 1.08 (1.02-1.14)               | 1.02 (0.90-1.13) | 0.99 (0.90-1.09) |
| Hypertension                 | 1.12 (1.05-1.22)               | 1.07 (0.93-1.25) | 0.89 (0.74-1.03) |
| Hypertension, Pulmonary      | 1.12 (1.00-1.25)               | 1.09 (0.88-1.42) | 1.03 (0.79-1.03) |
| Inflammatory Bowel Disease   | 1.10 (1.02-1.19)               | 0.97 (0.82-1.10) | 1.09 (0.93-1.29) |
| Knee arthroplasty            | 1.12 (1.02-1.25)               | 1.06 (0.86-1.35) | 0.99 (0.79-1.26) |
| Migraine                     | 2.08 (0.97-1.16)               | 1.02 (0.87-1.18) | 1.04 (0.85-1.31) |
| Osteoarthritis               | 1.07 (0.94-1.17)               | 1.05 (0.85-1.29) | 1.00 (0.80-1.24) |
| Osteoporosis                 | 1.12 (1.02-1.24)               | 1.09 (0.91-1.38) | 1.03 (0.82-1.34) |
| Parkinson's Disease (all)    | 1.10 (0.99-1.20)               | 1.04 (0.83-1.26) | 0.94 (0.71-1.16) |
| Psoriasis                    | 1.10 (0.97-1.24)               | 1.07 (0.85-1.37) | 1.00 (0.75-1.33) |
| Psoriatic arthropathy        | 1.11 (0.99-1.25)               | 1.08 (0.88-1.34) | 0.96 (0.73-1.21) |
| Pulmonary fibrosis           | 1.09 (0.99-1.19)               | 1.18 (0.97-1.69) | 1.08 (0.87-1.44) |
| Restless legs syndrome       | 1.09 (0.98-1.19)               | 0.98 (0.76-1.15) | 1.03 (0.82-1.31) |
| Rheumatoid arthritis         | 1.28 (1.08-1.30)               | 1.08 (0.95-1.23) | 1.09 (0.93-1.30) |
| Thromboembolism              | 1.10 (1.03-1.17)               | 1.07 (0.94-1.22) | 0.96 (0.85-1.07) |

1 - Trial level models adjusted for age, sex and comorbidity count. 2 - Trial level models adjusted for age and sex.

**Table S2 Odds ratio for association between attrition and age, sex and comorbidity count by ATC drug class**

| <b>ATC drug class</b> | <b>Drug type</b>                                     | <b>Comorbidity count<sup>1</sup></b> | <b>Age<sup>2</sup></b> | <b>Sex<sup>2</sup></b> |
|-----------------------|--|--------------------------------------|------------------------|------------------------|
| A10BH                 | Dipeptidyl peptidase 4 (DPP-4) inhibitors            | 1.08 (0.98-1.16)                     | 1.08 (0.94-1.29)       | 0.95 (0.81-1.09)       |
| A10BJ                 | Glucagon-like peptide-1 receptor analogues           | 1.09 (1.00-1.16)                     | 1.02 (0.86-1.14)       | 1.07 (0.92-1.28)       |
| A10BK                 | Sodium-glucose co-transporter 2 inhibitors           | 1.09 (1.03-1.16)                     | 1.02 (0.89-1.15)       | 0.96 (0.84-1.09)       |
| B01AB                 | Heparin group  | 1.10 (0.97-1.24)                     | 1.06 (0.90-1.29)       | 1.06 (0.90-1.26)       |
| B01AE                 | Direct thrombin inhibitors                           | 1.11 (1.05-1.17)                     | 1.08 (0.96-1.23)       | 0.92 (0.80-1.04)       |
| C09CA                 | Angiotensin II antagonists                           | 1.12 (1.04-1.21)                     | 1.06 (0.94-1.22)       | 0.90 (0.76-1.03)       |
| G04BE                 | Drugs used in erectile dysfunction                   | 1.13 (1.04-1.26)                     | 1.10 (0.96-1.35)       | 1.03 (0.80-1.36)       |
| H05AA                 | Parathyroid hormones and analogues                   | 1.11 (1.01-1.23)                     | 1.07 (0.91-1.31)       | 1.04 (0.83-1.33)       |
| L01XE                 | Protein kinase inhibitors                            | 1.09 (0.99-1.18)                     | 1.12 (0.95-1.49)       | 1.08 (0.88-1.40)       |
| L04AA                 | Selective immunosuppressants                         | 1.11 (1.01-1.22)                     | 1.03 (0.84-1.19)       | 1.07 (0.89-1.34)       |
| L04AB                 | Tumour necrosis factors alpha inhibitors             | 1.11 (1.04-1.19)                     | 1.01 (0.88-1.12)       | 1.06 (0.91-1.25)       |
| L04AC                 | Interleukin inhibitors                               | 1.17 (1.08-1.31)                     | 1.06 (0.95-1.21)       | 1.07 (0.92-1.27)       |
| M01AC                 | Oxicams  | 1.07 (0.93-1.17)                     | 1.05 (0.88-1.25)       | 1.01 (0.82-1.25)       |
| N03AX                 | Other antiepileptics                                 | 1.08 (0.97-1.16)                     | 1.03 (0.87-1.16)       | 1.04 (0.85-1.30)       |
| N04BC                 | Dopamine agonists                                    | 1.08 (0.98-1.16)                     | 1.00 (0.81-1.13)       | 0.98 (0.81-1.17)       |
| N06DA                 | Anticholinesterase drugs                             | 1.09 (1.01-1.18)                     | 1.06 (0.91-1.26)       | 0.89 (0.73-1.06)       |
| R03DX                 | Other systemic drugs for obstructive airway diseases | 1.08 (0.95-1.17)                     | 0.99 (0.76-1.13)       | 1.00 (0.78-1.26)       |

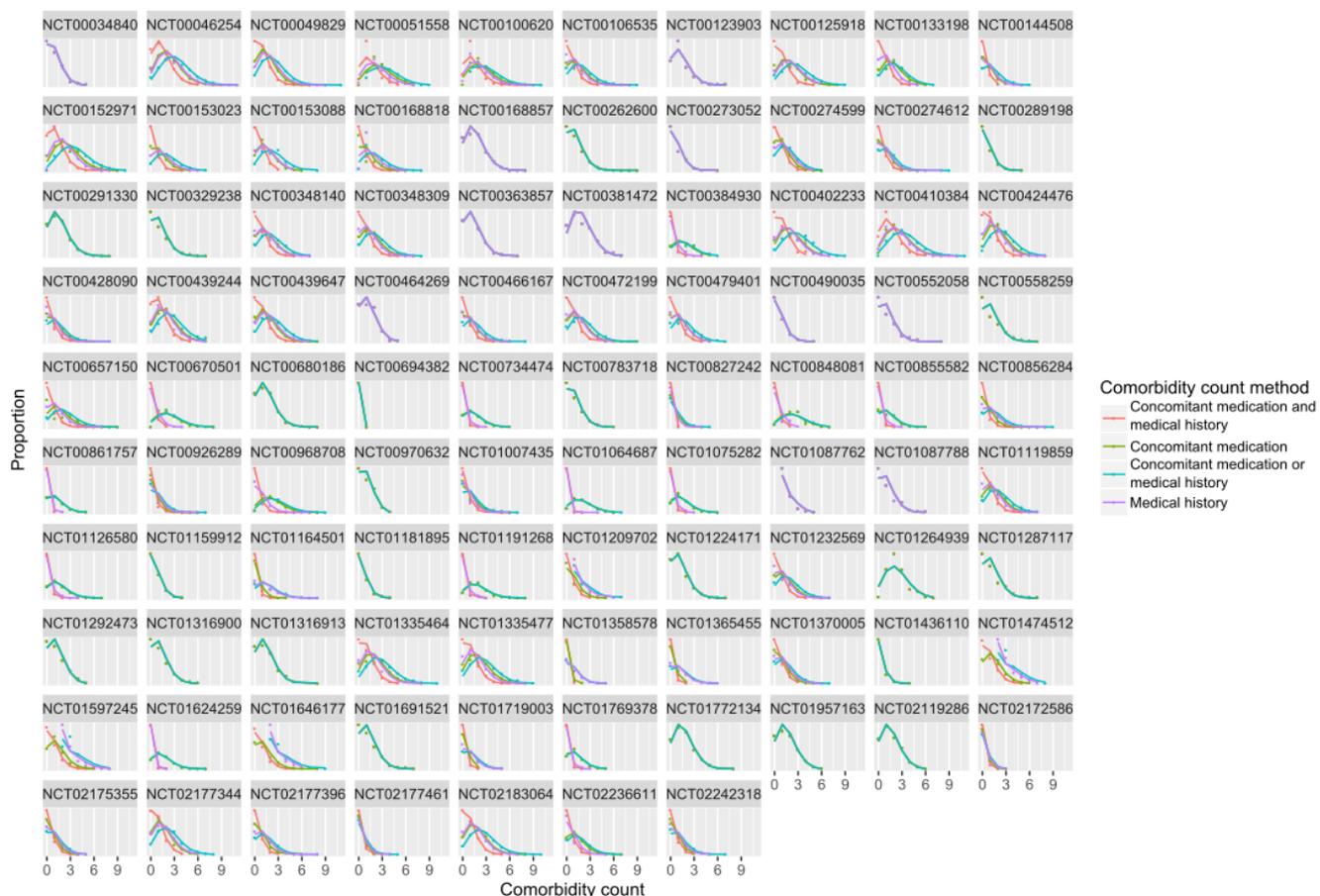
1 - Trial level models adjusted for age, sex and comorbidity count. 2 - Trial level models adjusted for age and sex.

**Table S3 Sensitivity analysis with wider priors on all of the variance components (trial, drug class and index condition)**

|  | <b>Original priors<sup>1</sup></b>          | <b>Wider priors<sup>2</sup></b>             |
|--|---|---|
| <b>Partial pooling across index condition and drug class</b> | N=20 index conditions;<br>N=17 drug classes | N=20 index conditions;<br>N=17 drug classes |
| Intercept (SD)   | 0.097 (0.026)                               | 0.096 (0.025)                               |
| SD trial (SD)  | 0.091 (0.018)                               | 0.092 (0.018)                               |
| SD index condition (SD)                                      | 0.044 (0.027)                               | 0.043 (0.027)                               |
| SD drug class (SD)   | 0.036 (0.026)                               | 0.035 (0.026)                               |

1 - Trial level models adjusted for age, sex and comorbidity count using original priors. 2 - Trial level models adjusted for age, sex and comorbidity count using wider priors for variance components. SD: standard deviations (and their SDs) for within-group variation for trials, index conditions and drug classes on log-odds scale. See supplementary methods below for full description of models and selection of priors.

Figure S1



Distribution of comorbidity count for each trial. Points indicate raw proportions. Lines indicate expected proportion under Poisson. The y-axis has been deliberately omitted and the resolution kept low to maintain privacy. The purpose of this plot is to show that the comorbidity counts follow a broadly Poisson distribution as justification for presenting mean counts as summary statistics and to explain the modelling choices used in the analysis of the impact of increasing comorbidity on overall trial attrition.

# Detailed description of modelling

## Model fitting

For each trial, we fitted generalised linear models to individual-level participant data within each trial repository.

$$\text{inv. logit}(p_i) = \beta_0 + \beta_1 \cdot \text{age} + \beta_2 \cdot \text{sex} + \beta_3 \cdot \text{comorbidity}_i$$

Where  $p_i$  was the probability of attrition and the  $\beta$  parameters 0, 1, 2 and 3 were the coefficients for the intercept, age, sex and comorbidity count respectively.  $i$  indicates the individual within each trial. A model was fitted separately for each of the  $j$  trials (subscript not shown above for clarity.) For each trial, we exported the resultant model coefficients ( $\beta_0, \beta_1, \beta_2, \beta_3$ ) standard errors ( $se_{\beta_0}, se_{\beta_1}, se_{\beta_2}, se_{\beta_3}$ ), and variance-covariance matrices for subsequent meta-analysis. The variance-covariance matrices had been exported in case more complex models needed to be fitted. However, as there was no evidence of departure from linearity, these were not used in any of the subsequent analysis, but are available on our github repository in case they are useful for other researchers.

The following shows the meta analysis for comorbidity for the most complex model shown in Table 2. The analyses for age and sex were similar. The simpler models excluded, respectively, the condition related terms, drug class related terms or both. The model had a normal likelihood and linear predictor as follows:-

$$\beta_{3,j} \sim N(\theta_j, se_{\beta_{3,j}})$$
$$\theta_j = \alpha_0 + \text{trial}_j + \text{cond}_k + \text{drug}_l$$

$\alpha_0$  indicates the overall effect. Trial, cond and drug indicate the trial-level, condition-level and drug-level random effects for the  $j$ -trials,  $k$ -conditions and  $l$ -drug classes. Note that trials were nested within drug classes and conditions, but conditions and drug classes were cross-classified (some of the drugs used to treat inflammatory conditions were used across multiple conditions).

The variation in the comorbidity-attrition association at the trial, condition and drug class effects was assumed to be normally distributed (random effects):-

$$\text{trial}_j \sim N(\mu_{\text{trial}}, \sigma_{\text{trial}})$$

$$\text{cond}_k \sim N(\mu_{\text{cond}}, \sigma_{\text{cond}})$$

$$\text{drug}_l \sim N(\mu_{\text{drug}}, \sigma_{\text{drug}})$$

The prior for the overall intercept  $\alpha$  was t-distributed with 3 degrees of freedom:-

$$\alpha_0 \sim \text{student-t}(df = 3, \text{mean} = 0, \text{sd} = 100)$$

For the trial, condition and drug-level random effects, the priors for mean and standard deviation were t-distributed and half-t distributed respectively:-

$$\mu_{\text{trial}} \sim \text{student-t}(df = 3, \text{mean} = 0, \text{sd} = 2.5)$$

$$\mu_{\text{cond}} \sim \text{student-t}(df = 3, \text{mean} = 0, \text{sd} = 2.5)$$

$$\mu_{\text{drug}} \sim \text{student-t}(df = 3, \text{mean} = 0, \text{sd} = 2.5)$$

$$\sigma_{trial} \sim \text{half-student-t}(df = 3, mean = 0, sd = 2.5)$$

$$\sigma_{cond} \sim \text{half-student-t}(df = 3, mean = 0, sd = 2.5)$$

$$\sigma_{drug} \sim \text{half-student-t}(df = 3, mean = 0, sd = 2.5)$$

As a sensitivity analysis of the effect of the selected priors, the models were re-run using wider priors with half-t distribution as follows:

$$\sigma_{trial} \sim \text{half-student-t}(df = 3, mean = 0, sd = 10)$$

$$\sigma_{cond} \sim \text{half-student-t}(df = 3, mean = 0, sd = 10)$$

$$\sigma_{drug} \sim \text{half-student-t}(df = 3, mean = 0, sd = 10)$$

$e^{\alpha_0}$  was the overall estimate for the odds ratio (OR) across all drugs and conditions. The condition-level and drug-level odds ratios were estimated as follows using samples from the posterior:-

$$OR_{cond_k} = e^{\alpha_0 + cond_k}$$

$$OR_{drug_l} = e^{\alpha_0 + drug_l}$$

For all parameters the mean of the un-transformed distribution was obtained and this was subsequently exponentiated to obtain the point estimate for the OR. The credible intervals were obtained as the 2.5th and 97.5th, 10th and 90th and 25th and 75th percentiles.

## Application to notional trial-level data

The point estimate for the OR was applied to simulated trial data as follows:-

1. For a set of 5 mean comorbidity counts (0.5, 0.75, 1, 1.25, 2) the probability mass function was used to calculate the proportion of participants with 0, 1, 2 ... 100 comorbidities.
2. For each of the above 5 sets of comorbidity counts, the risk of attrition in individuals with zero comorbidities was set to 0.05, 0.15, 0.25, 0.35 and 0.45, giving a total of 25 notional trials (5 mean comorbidity counts X 5 baseline risks of attrition).
3. Each baseline attrition risk was transformed to obtain the baseline odds of attrition.
4. For each comorbidity count (0 to 100), we estimated the odds ratio as  $OR^{comorbidity\ count}$ . For example,  $1.1^0 = 1$ ,  $1.1^1 = 1.1$ ,  $1.1^2 = 1.21$  for comorbidity counts 0, 1 and 2.
5. For each comorbidity count at each level of baseline risk, we estimated the comorbidity-count specific odds of attrition by multiplying the comorbidity-count specific odds ratio (step 4) by the baseline odds of attrition (step 3).
6. We transformed each of the odds from step 5 to obtain the comorbidity-count/baseline risk specific risk of attrition.
7. For each of the 25 notional trials, we calculated a weighted mean of the comorbidity-count specific risk of attrition, where the weights were the probabilities obtained in step 1. This is the trial-level risk of attrition.

The assumptions of this approach are as follows:-

1. Trial comorbidity is Poisson distributed.
2. The odds ratios are applicable to a set of notional trials with these levels for comorbidity count and baseline risk of attrition.