## Supplementary appendix

Centre for Mathematical Modelling of Infectious Diseases (CMMID) COVID-19 working group ..... 1
Table 1. List of conditions in GBD 2017 with potential to increase the risk of severe COVID-19 illness ..... 2
Methods used to calculate the proportion of individuals with at least one underlying condition ..... 3
Proportion with at least one underlying condition relevant to severe COVID-19 disease ..... 3
Adjustment for multimorbidity ..... 3
Figure 1. Empirical estimates of the ratio $r$ (left panel) and the multimorbidity fraction among those with at least one underlying health condition relevant to COVID-19 (right panel) from cross-sectional studies in Scotland and Southern China ..... 4
Methods to estimate numbers of individuals at high risk ..... 5
Infection hospitalisation ratios ..... 5
Adjustments to IHRs .....  .5
Relative risks for hospital admission .....  .5
Table 2. Summary of evidence on risk factors for COVID-19 hospital admission: used in a scenario to estimate of the number of individuals at high risk of severe COVID disease .....  .7
Sensitivity analysis for estimates of the number of individuals at high risk ..... 11
Table 3. Population in millions (\%) at high risk by region: base case and alternative scenarios ..... 12
Table 4. Low estimates of the number of individuals in millions (\% of total population) at increased risk of severe COVID-19 illness by age, number of conditions, region and age threshold ..... 13
Table 5. High estimates of the number of individuals in millions (\% of total population) at increased risk of severe COVID-19 illness by age, number of conditions, region and choice of age threshold... 14
Figure 2. Percentage of global population at increased risk of severe COVID-19 disease (left panel) and \% change (right panel) when conditions are removed one at a time ..... 15
Figure 3. Alternative version of Figure 3 in main paper showing the age-standardised proportion of population at increased risk and high risk of severe COVID-19 by country and region ..... 16
References ..... 17

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Table 1. List of conditions in GBD 2017 with potential to increase the risk of severe COVID-19 illness

| \# | Category | Causes included in the Global Burden of Disease Study (GBD2017) |
| :---: | :---: | :---: |
| 1 | HIV/AIDS | HIV/AIDS - Drug-susceptible Tuberculosis; HIV/AIDS - Multidrug-resistant Tuberculosis without extensive drug resistance; HIV/AIDS - Extensively drug-resistant Tuberculosis; HIV/AIDS resulting in other diseases |
| 2 | Tuberculosis* | Drug-susceptible tuberculosis; Multidrug-resistant tuberculosis without extensive drug resistance; Extensively drug-resistant tuberculosis |
| 3 | Cancers with direct immune suppression | Hodgkin lymphoma; Non-Hodgkin lymphoma; Multiple myeloma; Acute lymphoid leukemia; Chronic lymphoid leukemia; Acute myeloid leukemia; Chronic myeloid leukemia; Other leukemia; Other malignant neoplasms; Myelodysplastic, myeloproliferative, and other hematopoietic neoplasms |
| 4 | Cancers with possible immune suppression (from treatment therapy) | Lip and oral cavity cancer; Nasopharynx cancer; Other pharynx cancer; Esophageal cancer; Stomach cancer; Colon and rectum cancer; Liver cancer due to hepatitis B; Liver cancer due to hepatitis B; Liver cancer due to hepatitis B; Liver cancer due to hepatitis C; Liver cancer due to alcohol use; Liver cancer due to NASH; Liver cancer due to other causes; Gallbladder and biliary tract cancer; Pancreatic cancer; Larynx cancer; Tracheal, bronchus, and lung cancer; Malignant skin melanoma; Breast cancer; Cervical cancer; Uterine cancer; Ovarian cancer; Prostate cancer; Testicular cancer; Kidney cancer; Bladder cancer; Brain and nervous system cancer; Thyroid cancer; Mesothelioma |
| 5 | Cardiovascular disease | Rheumatic heart disease; Ischemic heart disease; Ischemic stroke; Intracerebral haemorrhage; Subarachnoid haemorrhage; Hypertensive heart disease; Non-rheumatic calcific aortic valve disease; Non-rheumatic degenerative mitral valve disease; Other non-rheumatic valve diseases; Myocarditis; Alcoholic cardiomyopathy; Other cardiomyopathy; Atrial fibrillation and flutter; Aortic aneurysm; Peripheral artery disease; Endocarditis; Other cardiovascular and circulatory diseases; Congenital heart anomalies |
| 6 | Chronic respiratory disease | Chronic obstructive pulmonary disease; Silicosis; Asbestosis; Coal workers pneumoconiosis; Other pneumoconiosis; Asthma*; Interstitial lung disease and pulmonary sarcoidosis |
| 7 | Chronic liver disease | Cirrhosis and other chronic liver diseases due to hepatitis B; Cirrhosis and other chronic liver diseases due to hepatitis C; Cirrhosis and other chronic liver diseases due to alcohol use; Cirrhosis and other chronic liver diseases due to other causes |
| 8 | Diabetes | Diabetes mellitus type 1; Diabetes mellitus type 2 |
| 9 | Chronic kidney disease | Chronic kidney disease due to diabetes mellitus type 1; Chronic kidney disease due to diabetes mellitus type 2; Chronic kidney disease due to hypertension; Chronic kidney disease due to glomerulonephritis; Chronic kidney disease due to other and unspecified causes |
| 10 | Chronic neurological disorders | Alzheimer's disease and other dementias; Parkinson's disease; Multiple sclerosis; Motor neuron disease; Other neurological disorders; Idiopathic developmental intellectual disability; Down syndrome; Neural tube defects |
| 11 | Sickle cell disorders | Sickle cell disorders |

* we excluded latent tuberculosis and adjusted asthma to better reflect BTS steps 4+.


## Methods used to calculate the proportion of individuals with at least one underlying condition

Note: Some of the text from the main paper is repeated here for convenience.

## Proportion with at least one underlying condition relevant to severe COVID-19 disease

The GBD study provides prevalence estimates for each disease category separately, but not what we needed, which was the prevalence of people in at least 1 of these categories. Diseases may cluster, for example if they are causally related. To deal with this, we first calculated $e$, which is the expected proportion of individuals with at least one condition assuming no clustering and that various prevalences are independent (e.g. the fact that someone has diabetes does not affect their risk of getting cancer) as 1 minus the probability of not having any of the conditions $\mathrm{c} 1, \mathrm{c} 2$, $\mathrm{c} 3 \ldots .$. i.e. $1-\left(1-\mathrm{p} \_\mathrm{c} 1\right) \mathrm{x}\left(1-\mathrm{p} \_\mathrm{c} 2\right) \mathrm{x}\left(1-\mathrm{p} \_\mathrm{c} 3\right) \ldots$.

We then estimated the proportion $P$, who have at least one underlying condition as $P=e \times r$, where $r$ is the ratio between the observed and expected percentage of individuals with at least one condition. We based $r$ on evidence from large cross-sectional multimorbidity studies in Scotland ${ }^{1}$ and Southern China. ${ }^{2}$

The ratio $r$ was broadly consistent by age, sex and study (see Figure 1 overleaf). For the analysis of both males and females combined, the mean of all age-specific values of $r$ was 0.92 (range 0.86 to 0.99 ) in Scotland and 0.92 (range ( $0.75-1.15$ ) in China. When extrapolating this value to other countries, we used a transparent ratio of 0.9 for all age groups and varied this this between 0.7 (low estimates) and 1.0 (high estimates). The resulting national estimates of $P$ were constrained to be no less than each country's single most prevalent condition. We conducted sensitivity analyses to explore the impact on results of using the observed age-specific values of $r$ rather than the same value for all ages, but this had a very small impact on the share of the population estimated to be at increased risk i.e. the share of the population at increased risk changed from $22.5 \%$ to $22.8 \%$ with adjustment for age.

## Adjustment for multimorbidity

In addition to providing estimates for $r$, the studies in Scotland and Southern China were also used to calculate the multimorbidity fraction i.e. the proportion of individuals with multiple (two or more) underlying conditions among those with at least one, by age group and sex. All analyses were done using disease categories that matched as closely as possible to the COVID-19-relevant categories defined in our analysis. In both studies this included: CVD (defined as the presence of one or more of coronary heart disease, hypertension, cerebrovascular disease, peripheral arterial disease, heart failure, or atrial fibrillation); chronic neurological disease (defined as one or more of dementia, multiple sclerosis and Parkinson's disease); and CRD (defined as one or both of chronic obstructive pulmonary disease and bronchiectasis). Other COVID-related conditions listed in the main analysis were counted separately. The GBD provide separate estimates for hypertensive heart disease and CKD due to hypertension, but it was not possible to make this distinction in the multimorbidity datasets, so all hypertension was included in the CVD category.

We calculated pooled estimates of the multimorbidity fraction by age and sex and extrapolated these pooled estimates to all countries included in the analysis (see Figure 1 overleaf).

Figure 1. Empirical estimates of the ratio $r$ (left panel) and the multimorbidity fraction among those with at least one underlying health condition relevant to COVID-19 (right panel) from cross-sectional studies in Scotland and Southern China

The top row shows results for females and males combined. The middle row shows results for females only and the bottom row shows results for males only.

The left panel/column shows the ratio between the observed and expected $\%$ of individuals with at least one condition by age. Expected estimates were calculated by assuming the prevalences of COVID-19 underling conditions are independent (e.g. the fact that someone has diabetes does not affect their risk of getting cancer) as 1 minus the probability of not having any of the conditions c 1 , c 2 , $\mathrm{c} 3 \ldots .$. i.e. $1-\left(1-\mathrm{p} \_\mathrm{c} 1\right) \times\left(1-\mathrm{p} \_\mathrm{c} 2\right) \mathrm{x}\left(1-\mathrm{p} \_\mathrm{c} 3\right) \ldots$. This was then compared to the observed value of the \% of individuals with at least one condition based on the same dataset (either Scotland or Southern China). The ratios between expected and observed are shown on the left panel/column below. Both studies indicate that the expected value based on the assumption of independence would provide reasonable estimates of the observed value. In our main analysis we assumed the ratio was 0.9 but varied this between 0.7 and 1.0 in uncertainty analysis.

The right panel/column shows the proportion of those with at least one underlying condition relevant to COVID-19 with multimorbidity (two or more conditions). As expected, this percentage increases with age in both studies. The grey lines represent pooled estimates and $95 \%$ confidence intervals based on a $2^{\text {nd }}$ order polynomial model fitted to all data points. Pooled estimates were extrapolated to all countries included in the analysis by age and sex. The lower and upper CI values were used in our low and high estimates in the main paper.

## Empirical estimates of the ratio $r$ by age Empirical estimates of multimorbidity by age



## Methods to estimate numbers of individuals at high risk

## Infection hospitalisation ratios

To estimate the number of individuals at high risk (those that would require hospital admission if infected) we applied country-level UN estimates of the number of individuals alive in each 5-year age group ${ }^{3}$ to age-specific infection hospitalisation ratios (IHRs) recently estimated for mainland China by Verity et al. ${ }^{4}$ IHRs represent the proportion of people who are infected (asymptomatic or symptomatic) that would have symptoms severe enough to require hospital admission. The term 'require hospital admission' is consistent with the WHO definition for severe cases. ${ }^{5}$ Whether or not these individuals receive hospital care is beyond the scope of this analysis and will depend on the health system in the country concerned.

## Adjustments to IHRs

We made two adjustments to account for differences between IHRs in China and other countries. The first was designed to capture the effect on IHRs of national variations in prevalence mix. The second was to adjust for infections in given age groups being more severe in higher mortality settings.

1. Adjusting for underlying conditions. For each 5-year age group and sex, the prevalence rates for each underlying condition were multiplied by their respective relative risks (RRs) for hospitalisation. RRs of 3.0 were assumed for CKD, diabetes and CVD and of 1.5 for the eight other conditions. The RRs we used were informed by a rapid review of what is currently known about the strength of association between different variables and COVID-19 hospital admission (see table 2 below). The totals were then summed across all 11 conditions and added to the proportion of individuals without underlying conditions, to create a risk score for each 5-year age group. IHRs were then adjusted to account for the ratio of the risk score for China and the country of interest. For example, for males aged 55-59 years the risk score was 1.95 in China and 2.63 in Afghanistan, so the IHR for this age group was multiplied by a ratio of 1.35 (2.63/1.95). This adjustment assumes that for each underlying condition, the RR of hospital admission is the same for every country.
2. Adjusting for age-based frailty. Direct interpretation of the first adjustment relies on constant RRs of admission for each condition across countries. To address the likelihood of more severe infections at given ages in higher mortality settings we went further. For each 5-year age group and sex, we divided UN estimates of age-specific life expectancy ${ }^{3}$ for China by the equivalent estimate for the country of interest. This ratio (proxy for difference in age-based frailty) was then applied to the IHR for same age group. For example, the life expectancy for males at age 55 years is 22.8 and 19.1 years in China and Afghanistan, respectively, so the IHR for this age group was multiplied by a ratio of 1.19 (22.8/19.1) to generate the estimate for Afghanistan. Country-specific differences in life expectancy may be more extreme than country-specific differences in infection severity because life expectancy incorporates not only general health status, but also access to effective quality health care. Also, the extent to which infections will be more severe in higher mortality settings is still very uncertain. We therefore show results with and without this adjustment.

Thus, for males aged 55-59 years in Afghanistan, separate adjustments for underlying conditions and age-based frailty had the effect of increasing the IHR by 1.61 (IHR x $1.35 \times 1.19$ ), from $9 \%$ to $15 \%$.

## Relative risks for hospital admission

We searched PubMed ("Risk factors" AND "COVID-19") without language restrictions, from database inception until April 5, 2020, and identified 62 studies published between Feb 15, 2020 and March 20, 2020. We later ran a separate PubMed search ("Multivariable" AND "COVID-19") without language restrictions, from database inception until May 25, 2020.

Several studies are emerging on the risk of mortality among those already hospitalised ${ }^{6,7}$ but we restricted our analysis of RRs to studies that allowed comparison of hospitalised and non-hospitalised COVID-19 cases. This was because our analysis focused on the risk of severe disease (requiring hospital admission) rather than the risk of death. Three studies from the USA met these criteria. The first contained descriptive data on 6,637 COVID19 cases reported to the CDC as of $28^{\text {th }}$ March, 2020. For this study, we derived crude univariable RRs for each condition from the reported case counts. ${ }^{8}$ The second was a multivariable analysis of 4,103 COVID-19 cases in New York City, between $1^{\text {st }}$ March 2020 and $2^{\text {nd }}$ April 2020. ${ }^{9}$ The third was a multivariable study from Northern

California with 1,052 confirmed cases of COVID-19 captured between ${ }^{\text {st }}$ January and $8^{\text {th }}$ April, 2020. ${ }^{10}$ For conditions where evidence was weak or missing, we included studies that did not meet our initial inclusion criteria, either because they were not yet published (i.e. multivariable analyses in Italy ${ }^{11}$ and the $\mathrm{UK}^{12}$ ) or because many of the COVID-19 cases in the less severe group were hospitalised i.e. one published metaanalysis of 4 studies from China. ${ }^{13}$ Using the evidence from these studies, we summarised the strength of the association with hospital admission (low, moderate, high) and graded our confidence in the strength of that association (low, moderate high).

This rapid review aimed to summarise what is currently known about the strength of association between different variables and COVID-19 hospital admission, but this is unlikely to be exhaustive. We therefore chose to use a very simple stratification of risk for the 11 conditions (either $R R=3$ or $R R=1.5$ ) based on the range of RRs reported for these conditions. For three conditions (CKD, diabetes and CVD) we had moderate confidence that the strength of the association would be high based on univariable and multivariable analysis, so assumed a RR of 3.0. For the eight other conditions we assumed a RR of 1.5 because there was insufficient evidence of a significant independent association with hospital admission from multivariable analyses.

RR values for each condition were varied one at a time (assuming low and high values of 1 and 10 respectively) to assess the impact of these changes on the total population at high risk (p 11).

Table 2. Summary of evidence on risk factors for COVID-19 hospital admission: used in a scenario to estimate of the number of individuals at high risk of severe COVID disease

| Variable | What is known about risk factors for COVID-19 hospital admission? <br> (Strength of association is based on evidence restricted to studies that include a control group of COVID-19 patients that were not severe enough to be admitted to hospital) | Strength of association with admission (low, moderate, high) | Confidence about strength of association (low, moderate, high) |
| :---: | :---: | :---: | :---: |

Underlying conditions listed in guidelines with available data by age, sex and country

| Cardiovascular disease | Crude RR $=5.39$ ( $95 \%$ CI $4.63-6.27$ ) based on $25 \%$ vs $5 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ). ${ }^{8}$ Heart failure was also a significant independent predictor (OR $=4.29$ [1-89-11.18]) in multivariable analysis in New York City ( $n=4103$ ). ${ }^{9}$ Congestive heart failure was also associated with increased odds of hospital admission $(\mathrm{OR}=3.3)$ based on multivariable analysis from Northern California ( $\mathrm{n}=1,052$ ). ${ }^{10}$ | High | Moderate |
| :---: | :---: | :---: | :---: |
| Chronic kidney disease | Crude RR $=10.19$ ( $95 \%$ CI $7.46-13.93$ ) based on $10 \%$ vs $1 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ). ${ }^{8}$ CKD was also a significant independent predictor ( $\mathrm{OR}=3.07$ [1-78-5.52]) in multivariable analysis in New York City ( $\mathrm{n}=4103$ ). ${ }^{9}$ | High | Moderate |
| Diabetes | Crude RR $=4.15$ ( $95 \%$ CI $3.63-4.74$ ) based on $27 \%$ vs $6 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ). ${ }^{8}$ Diabetes was also a significant independent predictor ( $\mathrm{OR}=$ 2.81 [2.14-3.72]) in multivariable analysis in New York City ( $\mathrm{n}=4103$ ). ${ }^{9}$ Type 2 diabetes was also associated with increased odds of hospital admission $(\mathrm{OR}=2.2)$ based on multivariable analysis from Northern California ( $\mathrm{n}=1,052$ ).$^{10}$ | High | Moderate |
| Neurological disorders | Crude RR $=6.48$ ( $95 \%$ CI $3.61-11.64$ ) based on $2 \%$ vs $0 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ), but the number with neurological disorders was small in this study ( $n=49$ ), ${ }^{8}$ and results on the association with hospital admission were not reported in the multivariable analyses. Data from an unpublished multivariable analysis in Italy ( $\mathrm{n}=2143$ ) found a non-significant association between dementia and hospital admission a (HR = 1.2 [95\% CI 0.9-1.7]). ${ }^{11}$ | Moderate | Low |
| Chronic respiratory diseases | Crude RR $=2.33$ ( $95 \%$ CI $2.01-2.71$ ) based on $16 \%$ vs $7 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ). ${ }^{8}$ Not a significant independent predictor ( $\mathrm{OR}=1.33$ [0.96-1.84]) in multivariable analysis in New York City $(\mathrm{n}=4103) .{ }^{9}$ <br> However, both studies included all severities of 'asthma' which is likely to deflate the true odds associated with other chronic respiratory diseases. A meta-analysis of 4 studies from China estimated a pooled univariable OR of 2.46 (95\% CI $1.76-$ $3.44, \mathrm{I}$-squared $=0.0 \%, \mathrm{p}=0.611$ ) but included many nonsevere cases that were hospitalised. ${ }^{14}$ Data from an unpublished multivariable analysis in Italy ( $\mathrm{n}=2143$ ) found an association between chronic obstructive pulmonary disease (COPD) and hospital admission (HR $=1.9$ [95\% CI 1.4-2.5]). ${ }^{11}$ | Moderate | Low |
| Tuberculosis (active) | Not data is available, so we assume the same RR assumed for chronic respiratory diseases. | Insufficient data | Low |
| Chronic liver disease | Crude RR $=2.29$ ( $95 \%$ CI $1.22-4.31$ ) based on $1 \%$ vs $0 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ), but the number with chronic liver disease was small in this study ( $\mathrm{n}=40$ ). ${ }^{8}$ An unpublished study from South Korea found | Moderate | Low |
| Cancers with possible immunosuppres sion | Crude RR $=2.54(95 \%$ CI $1.98-3.25)$ based on $7 \%$ vs $3 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ) based on a generic category of immunosuppressed conditions. ${ }^{8}$ Not a significant independent predictor $(O R=1.24$ [0.81-1.93]) in multivariable analysis in New York City ( $\mathrm{n}=4103$ ) based on a general malignancy category. ${ }^{9}$ Data from an unpublished multivariable analysis in Italy ( $\mathrm{n}=2143$ ) found an association | Moderate | Low |


| Variable | What is known about risk factors for COVID-19 hospital admission? <br> (Strength of association is based on evidence restricted to studies that include a control group of COVID-19 patients that were not severe enough to be admitted to hospital) | Strength of association with admission (low, moderate, high) | Confidence about strength of association (low, moderate, high) |
| :---: | :---: | :---: | :---: |
|  | between cancer and hospital admission (HR $=1.4$ [95\% CI 1.11.7]). ${ }^{11}$ |  |  |
| Cancers with direct immunosuppres sion | Crude RR $=2.54$ (95\% CI 1.98-3.25) based on 7\% vs 3\% prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ) based on a generic category of immunosuppressed conditions. ${ }^{8}$ Data from an unpublished multivariable analysis in Italy ( $\mathrm{n}=2143$ ) found an association between cancer and hospital admission (HR $=1.4$ [95\% CI 1.1-1.7]). ${ }^{11}$ | Moderate | Low |
| HIV / AIDS | Crude RR $=2.54$ ( $95 \%$ CI $1.98-3.25$ ) based on $7 \%$ vs $3 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ) based on a generic category of immunosuppressed conditions. ${ }^{8}$ | Moderate | Low |
| Sickle cell disorders | Crude RR $=2.54$ (95\% CI 1.98-3.25) based on $7 \%$ vs $3 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ) based on a generic category of immunosuppressed conditions. ${ }^{8}$ | Moderate | Low |

## Important variables with available data in 188 countries

| Age | Current guidelines consider those aged 60+ (WHO), 65+ (US) and 70+ years (UK) to be at increased risk. Based on multivariable analysis in New York City ( $\mathrm{n}=4103$ ), ORs were 2.57 (2.06-3.20) aged 45-54yrs, 4.17 (3.35-5.20) aged 5564yrs, 10.91 aged $65-74 \mathrm{yrs}$ and 66.79 (44-73-102.62) aged $75+$ years, compared to the reference group aged 19-44 years. ${ }^{9}$ In this study the median age was 62 vs 41 years with and without admission. | High <br> (reflected in GBD prevalence data and IHRs, which both have a strong age effect) | High |
| :---: | :---: | :---: | :---: |
| Male gender | Gender is not included in current guidelines. In one multivariable analysis in New York City ( $\mathrm{n}=4103$ ) male gender was a significant independent predictor of hospitalisation (OR $=2.80,95 \%$ CI $2.38-3.30, \mathrm{p}<0.001) .{ }^{9}$ Males were twice as likely to be admitted to hospital as females ( $\mathrm{OR}=1.9, \mathrm{p}$ $=0.001$ ) based on multivariable analysis from Northern California ( $\mathrm{n}=1,052$ ) ${ }^{10}$ In a large study of hospitalised patients in the UK, $60 \%$ of hospitalised patients were male and this effect was seen across all ages $(\mathrm{n}=16,749) .{ }^{6}$ Males had a RR of 1.64 ( $95 \%$ CI $1.25-2.14, \mathrm{p}<0.001$ ) in a multivariable study that compared a cohort that did and did not have a test-positive hospital admission in the UK ( $\mathrm{n}=428,225$ ). ${ }^{12}$ Data from an unpublished multivariable analysis in Italy $(\mathrm{n}=1866)$ found an association between male gender and hospital admission (HR = 1.4 [95\% CI 1.2-1.6]). ${ }^{11}$ | High <br> (for simplicity we assumed males were twice as likely as females to be at high risk i.e. to require hospital admission if infected. | High |

Other possible variables not included in this analysis

|  | The prevalence of pregnancy was the same among COVID-19 <br> hospital admissions (6\%) and the wider community in a large <br> study in the UK (n=16749), ${ }^{6}$ but this is likely to remain in <br> current guidelines until more evidence emerges e.g. on the risks <br> at different stages of pregnancy. | Low | Low |
| :--- | :--- | :--- | :--- |
|  | GBD includes hypertensive heart disease within CVD and <br> CKD due to hypertension within CKD. Other forms of <br> hypertension are not included. Hypertension that is unrelated to <br> other underlying conditions is not included in current <br> guidelines, and it is unclear if prevalence data are available for |  |  |
| Hypertension <br> (excluding <br> hypertensive <br> heart disease <br> and CKD <br> caused by <br> hypertension) | 188 countries by age and sex. In one multivariable analysis in <br> New York City (n=4103), hypertension was not identified as a <br> significant independent predictor (OR $=1.23,95 \% ~ C I ~[0.97-$ <br> $1.57]$, p=0.094). ${ }^{9}$ Individuals taking blood pressure medication <br> had a RR of 1.40 (95\% CI $1.08-1.82, \mathrm{p}=0.01)$ in a <br> multivariable study that compared a cohort that did and did not <br> have a test-positive hospital admission in the UK | Low | Moderate |


| Variable | What is known about risk factors for COVID-19 hospital admission? <br> (Strength of association is based on evidence restricted to studies that include a control group of COVID-19 patients that were not severe enough to be admitted to hospital) | Strength of association with admission (low, moderate, high) | Confidence about strength of association (low, moderate, high) |
| :---: | :---: | :---: | :---: |
|  | ( $\mathrm{n}=428,225$ ). ${ }^{12}$ A meta-analysis of 4 studies from China estimated a pooled univariable OR of 2.36 ( $95 \%$ CI 1.46 3.83 , I-squared $=39.3 \%, \mathrm{p}=0.176$ ) but included many nonsevere cases that were hospitalised. ${ }^{14}$ Data from an unpublished multivariable analysis in Italy ( $\mathrm{n}=2143$ ) found an association between hypertension and hospital admission ( $\mathrm{HR}=1.4[95 \%$ CI 1.2-1.6]). ${ }^{11}$ |  |  |
| Current or former smoker | Included in WHO and US guidelines but not UK guidelines. Not a significant independent predictor ( $\mathrm{OR}=0.71,95 \%$ CI 0.57-0.87, $\mathrm{p}=0.001$ ) in multivariable analysis in New York City ( $n=4103$ ). ${ }^{9}$ However, current and former smokers were combined into a single category and former smokers appear to have a higher risk. For example, the crude RR for former smokers was 3.36 ( $95 \%$ CI $2.47-4.56$ ) based on $5 \%$ vs $2 \%$ prevalence with and without admission in the USA ( $\mathrm{n}=6637$ ) and ex-smokers were found to have worse survival after hospital admission in a large multivariable analysis in the UK ( $\mathrm{n}=17,425,445$ ) - the hazard ratio was 1.25 ( $95 \%$ CI 1.18-1.33) compared to those that never smoked. ${ }^{7}$ It is unclear whether prevalence data would be available for this category by 5-year age group, sex and country ( $\mathrm{n}=188$ ). | Moderate | Low |
| Ethnicity | Not included in current guidelines. Unclear if prevalence data are available for 188 countries by sex. In one multivariable analysis in New York City ( $n=4103$ ), other/multiracial race was identified as a significant independent predictor $(O R=1.99$, $95 \%$ CI [1.62-2.45], p<0.001) compared to white race, but African American and Asian race was not a significant predictor. ${ }^{9}$ The risk of hospital admission for African Americans was more than double that of non-Hispanic whites ( $\mathrm{OR}=2.7, \mathrm{p}=0.007$ ) in a multivariable study from Northern California ( $\mathrm{n}=1,052$ ). ${ }^{10}$ | Moderate | Moderate |
| BMI 30-39 | Not included in current guidelines. Significant independent predictor ( $\mathrm{OR}=4.26,95 \%$ CI $3.50-5.20, \mathrm{p}<0.001$ ) in multivariable analysis in New York City ( $\mathrm{n}=4103$ ). ${ }^{9}$ BMI of $30+$ had a RR of 1.97 ( $95 \%$ CI $1.46-2.65, \mathrm{p}<0.0001$ ) in a multivariable study that compared a cohort that did and did not have a test-positive hospital admission in the UK ( $\mathrm{n}=428,225$ ). ${ }^{12}$ Data from an unpublished multivariable analysis in Italy $(\mathrm{n}=2143)$ reported a $\mathrm{HR}=1.4$ [95\% CI 0.92.0]). ${ }^{11}$ Prevalence data are available for adults aged $18+$ years, but it is currently unclear if these are available for BMI 30+ by sex and 5-year age group. | High | Moderate |
| BMI 40+ | Included in US/UK guidelines but not WHO guidelines. Significant independent predictor ( $\mathrm{OR}=6.20,95 \%$ CI 4.21$9.25, \mathrm{p}<0.001$ ) in multivariable analysis in New York City ( $\mathrm{n}=4103$ ). ${ }^{9}$ Prevalence data are available for adults aged $18+$ years, but it is currently unclear if these are available for BMI $40+$ by sex and 5-year age group. | High | Moderate |
| Organ donor recipients | Included in current guidelines but unclear if prevalence data are available for 188 countries by 5-year age and sex. No studies have assessed the association with hospital admission, but it was found to be an important predictor of survival after hospital admission in a large multivariable analysis in the UK ( $\mathrm{n}=17,425,445$ ) - the hazard ratio was 4.27 ( $95 \%$ CI 3.205.70). ${ }^{7}$ | Insufficient data | Insufficient data |
| Laboratory markers | Not included in current guidelines. Insufficient data on the prevalence of markers in the general community (e.g. C reactive protein $>200 \mathrm{mg} / \mathrm{L}$, d-dimer $>500 \mathrm{ng} / \mathrm{mL}$, etc.) so insufficient evidence on risk of hospitalisation. Vitamin D status is emerging as a priority for research and may provide | Insufficient data | Insufficient data |


| Variable | What is known about risk factors for COVID-19 hospital admission? <br> (Strength of association is based on evidence restricted to studies that include a control group of COVID-19 patients that were not severe enough to be admitted to hospital) | Strength of association with admission (low, moderate, high) | Confidence about strength of association (low, moderate, high) |
| :---: | :---: | :---: | :---: |
|  | one explanation for the disproportionate number of Black, Asian and Minority Ethnic groups (BAME) experiencing severe symptoms in northern latitudes. ${ }^{15}$ |  |  |
| Child malnutrition | Not included in current guidelines. Prevalence data are available but there is insufficient evidence of risk associated with hospital admission. | Insufficient data | Insufficient data |
| Malaria | Not included in current guidelines. Prevalence data are available but there is insufficient evidence of risk associated with hospital admission. | Insufficient data | Insufficient data |
| Deprivation | Not included in current guidelines. Unclear if prevalence data are available for 188 countries. Insufficient evidence of risk associated with hospital admission, but was found to be an important predictor of survival after hospital admission in a large multivariable analysis in the UK $(\mathrm{n}=17,425,445)$ - the hazard ratio was 1.75 ( $95 \%$ CI 1.60-1.91) in the most deprived quintile compared to the least deprived. ${ }^{7}$ | Insufficient data | Insufficient data |
| Crowded housing | Not included in current guidelines. Unclear if prevalence data are available for 188 countries by age and sex. Insufficient evidence of risk associated with hospital admission. | Insufficient data | Insufficient data |
| Health and social care workers | Not included in current guidelines. Unclear if prevalence data are available for 188 countries by age and sex. Likely to be an important target group for a future vaccine due to increased risk of transmission. | Insufficient data | Insufficient data |
| Residents of care homes and other facilities | Included in some guidelines. Unclear if prevalence data are available for 188 countries. Residential care homes are less prevalent in lower income settings but are likely to be an important target group for a future vaccine in high income settings due to increased risk of transmission. | Insufficient data | Insufficient data |

## Sensitivity analysis for estimates of the number of individuals at high risk

Table 3 (overleaf) summarises the share of the population estimated to be at high risk (those that would require hospital admission if infected) for the base case scenario and for a range of alternative scenarios.

The following assumptions were varied:

1. Low and high credible intervals of IHRs - scenarios based on the low and high credible interval values of the IHRs reported in Verity et al, ${ }^{16}$ were influential. The global population at high risk (4.5\%) decreased to $2.7 \%$ with low IHRs and increased to $9.1 \%$ with high IHRs;
2. Adjustments for underlying conditions - removing this adjustment reduced the global population at high risk from $4.5 \%$ to $4.0 \%$ (and from $3.1 \%$ to $2.7 \%$ in Africa). In some countries removing this adjustment had a more substantial impact, due to very high prevalence of specific conditions relative to the same conditions in China e.g. diabetes in Fiji, HIV/AIDs in Swaziland;
3. Adjustments for age-based frailty - as expected, removing the age-based frailty adjustment resulted in a lower population at high risk in Africa (from $3.1 \%$ to $2.6 \%$ ) and a higher population at high risk in Europe and other high-income settings;
4. Altering the maximum proportion of the population infected - our central estimates of the numbers at high risk of severe COVID-19 disease assume there is no theoretical maximum to the proportion of the population that could ever be infected. Thus, the total number of individuals at high risk is calculated by multiplying the total population in each age group by the IHR for the same age group. However, as our understanding of COVID-19 transmission dynamics improves, empirical data may begin to emerge on the scale and duration of immunity acquired from natural infections, and on the proportion of the population that could ever be infected given widespread community transmission. If this is lower than our current assumption of $100 \%$, then fewer individuals would be at high risk using our method. When we varied this value, the global population at high risk decreased from $4.5 \%$ assuming all individuals could ever be infected to $4.0 \%, 3.6 \%, 3.1 \%$ and $2.7 \%$ when assuming this value could not exceed $90 \%$, $80 \%, 70 \%$ and $60 \%$, respectively; and,
5. Altering the RRs for specific conditions - changing the RR values for each condition one at a time (assuming low and high values of 1 and 10 respectively) had a limited impact on the total population at high risk (<5\% increase/decrease). Results were most sensitive to CVD, CKD, diabetes and liver disease (due to its higher prevalence in China relative to many other settings). Increasing the RR for HIV/AIDS from 1.5 to 10.0 was influential in Africa and increased the share of the population at high risk from $3.1 \%$ ( 42 million) to $3.7 \%$ ( 49 million).

Table 3. Population in millions (\%) at high risk by region: base case and alternative scenarios

|  | Africa | Asia | Europe | Latin <br> America <br> and the <br> Caribbean | Northern <br> America | Oceania |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |

Table 4. Low estimates of the number of individuals in millions (\% of total population) at increased risk of severe COVID-19 illness by age, number of conditions, region and age threshold
Assumptions:

- Population estimates:
- Disease prevalence:
- Ratio (observed vs expected $\%$ with $1+$ conditions): 0.7
- Multimorbidity among those with $1+$ conditions:

Lower 95\% CI
Lower 95\% CI

Lower 95\% CI

|  | Africa | Asia | Europe | Latin <br> America and the Caribbean | Northern America | Oceania | Global |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| No conditions |  |  |  |  |  |  |  |
| <15 years | 467 (40) | 948 (23) | 109 (16) | 137 (24) | 59 (18) | 9 (23) | 1729 (25) |
| 15-49 years | 508 (43) | 1910 (46) | 272 (40) | 276 (48) | 139 (42) | 17 (43) | 3122 (45) |
| 50-54 years | 27 (2) | 195 (5) | 36 (5) | 24 (4) | 15 (5) | 2 (4) | 299 (4) |
| 55-59 years | 20 (2) | 152 (4) | 34 (5) | 20 (3) | 15 (4) | 2 (4) | 242 (3) |
| 60-64 years | 14 (1) | 113 (3) | 28 (4) | 15 (3) | 12 (4) | 1 (3) | 184 (3) |
| 65-69 years | 9 (1) | 87 (2) | 22 (3) | 11 (2) | 9 (3) | 1 (2) | 139 (2) |
| 70+ years | 11 (1) | 109 (3) | 40 (6) | 15 (3) | 15 (5) | 2 (4) | 192 (3) |
| All ages | 1056 (90) | 3513 (85) | 542 (79) | 499 (86) | 265 (80) | 32 (85) | 5906 (85) |
| One condition only |  |  |  |  |  |  |  |
| <15 years | 8 (1) | 21 (1) | 1 (0) | 1 (0) | 0 (0) | 0 (0) | 31 (0) |
| 15-49 years | 59 (5) | 215 (5) | 29 (4) | 26 (5) | 11 (3) | 2 (4) | 342 (5) |
| 50-54 years | 9 (1) | 52 (1) | 9 (1) | 7 (1) | 4 (1) | 0 (1) | 81 (1) |
| 55-59 years | 8 (1) | 50 (1) | 11 (2) | 7 (1) | 6 (2) | 0 (1) | 82 (1) |
| 60-64 years | 7 (1) | 45 (1) | 12 (2) | 6 (1) | 6 (2) | 0 (1) | 77 (1) |
| 65-69 years | 6 (0) | 41 (1) | 12 (2) | 5 (1) | 6 (2) | 0 (1) | 71 (1) |
| 70+ years | 8 (1) | 67 (2) | 30 (4) | 11 (2) | 13 (4) | 1 (3) | 130 (2) |
| All ages | 104 (9) | 491 (12) | 105 (15) | 63 (11) | 47 (14) | 4 (12) | 814 (12) |
| Multiple (two or more) conditions |  |  |  |  |  |  |  |
| <15 years | 0 (0) | 0 (0) | 0 (0) | 0 (0) | 0 (0) | 0 (0) | 0 (0) |
| 15-49 years | 5 (0) | 20 (0) | 3 (0) | 3 (0) | 1 (0) | 0 (0) | 32 (0) |
| 50-54 years | 2 (0) | 11 (0) | 2 (0) | 1 (0) | 1 (0) | 0 (0) | 17 (0) |
| 55-59 years | 2 (0) | 14 (0) | 3 (0) | 2 (0) | 2 (0) | 0 (0) | 22 (0) |
| 60-64 years | 2 (0) | 16 (0) | 4 (1) | 2 (0) | 2 (1) | 0 (0) | 27 (0) |
| 65-69 years | 2 (0) | 18 (0) | 5 (1) | 2 (0) | 3 (1) | 0 (0) | 31 (0) |
| 70+ years | 5 (0) | 45 (1) | 22 (3) | 7 (1) | 9 (3) | 1 (2) | 89 (1) |
| All ages | 19 (2) | 124 (3) | 39 (6) | 18 (3) | 18 (5) | 1 (4) | 218 (3) |
| At increased risk of severe COVID-19 illness |  |  |  |  |  |  |  |
| At least one condition | 123 (10) | 614 (15) | 144 (21) | 81 (14) | 64 (20) | 6 (15) | 1032 (15) |
| Plus others older than: |  |  |  |  |  |  |  |
| 50 years | 204 (17) | 1269 (31) | 305 (44) | 166 (29) | 131 (40) | 13 (33) | 2088 (30) |
| 55 years | 178 (15) | 1074 (26) | 268 (39) | 142 (24) | 116 (35) | 11 (29) | 1789 (26) |
| 60 years | 157 (13) | 923 (22) | 234 (34) | 122 (21) | 101 (31) | 10 (25) | 1547 (22) |
| 65 years | 136 (12) | 722 (17) | 187 (27) | 97 (17) | 83 (25) | 7 (19) | 1233 (18) |
| 70 years | 131 (11) | 681 (17) | 175 (25) | 92 (16) | 77 (24) | 7 (18) | 1163 (17) |

Table 5. High estimates of the number of individuals in millions (\% of total population) at increased risk of severe COVID-19 illness by age, number of conditions, region and choice of age threshold
Assumptions:

- Population estimates:
- Disease prevalence:
- Ratio (observed vs expected $\%$ with $1+$ conditions): 1.0
- Multimorbidity among those with $1+$ conditions:

Upper 95\% CI
Upper 95\% CI

Upper 95\% CI

|  | Africa | Asia | Europe | Latin <br> America and the Caribbean | Northern America | Oceania | Global |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| No conditions |  |  |  |  |  |  |  |
| <15 years | 577 (38) | 1140 (22) | 126 (16) | 169 (23) | 72 (18) | 10 (23) | 2093 (24) |
| 15-49 years | 560 (37) | 2059 (40) | 279 (34) | 308 (43) | 156 (38) | 17 (38) | 3378 (39) |
| 50-54 years | 23 (2) | 175 (3) | 31 (4) | 22 (3) | 14 (3) | 1 (3) | 267 (3) |
| 55-59 years | 16 (1) | 126 (2) | 26 (3) | 16 (2) | 12 (3) | 1 (3) | 197 (2) |
| 60-64 years | 10 (1) | 83 (2) | 18 (2) | 11 (1) | 8 (2) | 1 (2) | 131 (2) |
| 65-69 years | 5 (0) | 56 (1) | 12 (1) | 6 (1) | 5 (1) | 1 (1) | 85 (1) |
| 70+ years | 4 (0) | 48 (1) | 12 (1) | 6 (1) | 4 (1) | 1 (1) | 73 (1) |
| All ages | 1193 (80) | 3686 (72) | 504 (62) | 538 (74) | 271 (66) | 32 (71) | 6224 (72) |
| One condition only |  |  |  |  |  |  |  |
| <15 years | 26 (2) | 61 (1) | 4 (0) | 4 (1) | 2 (0) | 0 (1) | 96 (1) |
| 15-49 years | 139 (9) | 509 (10) | 66 (8) | 61 (8) | 27 (7) | 4 (8) | 807 (9) |
| 50-54 years | 18 (1) | 108 (2) | 18 (2) | 14 (2) | 9 (2) | 1 (2) | 167 (2) |
| 55-59 years | 16 (1) | 100 (2) | 22 (3) | 13 (2) | 11 (3) | 1 (2) | 162 (2) |
| 60-64 years | 13 (1) | 88 (2) | 23 (3) | 12 (2) | 12 (3) | 1 (2) | 148 (2) |
| 65-69 years | 10 (1) | 77 (1) | 21 (3) | 10 (1) | 11 (3) | 1 (2) | 130 (2) |
| 70+ years | 13 (1) | 107 (2) | 43 (5) | 17 (2) | 20 (5) | 2 (4) | 201 (2) |
| All ages | 236 (16) | 1050 (20) | 197 (24) | 131 (18) | 90 (22) | 9 (20) | 1712 (20) |
| Multiple (two or more) conditions |  |  |  |  |  |  |  |
| $<15$ years | 3 (0) | 8 (0) | 1 (0) | 1 (0) | 0 (0) | 0 (0) | 13 (0) |
| 15-49 years | 26 (2) | 102 (2) | 14 (2) | 12 (2) | 6 (1) | 1 (2) | 161 (2) |
| 50-54 years | 6 (0) | 38 (1) | 6 (1) | 5 (1) | 3 (1) | 0 (1) | 58 (1) |
| 55-59 years | 7 (0) | 42 (1) | 9 (1) | 6 (1) | 4 (1) | 0 (1) | 68 (1) |
| 60-64 years | 7 (0) | 45 (1) | 12 (1) | 6 (1) | 6 (1) | 0 (1) | 75 (1) |
| 65-69 years | 7 (0) | 48 (1) | 13 (2) | 6 (1) | 7 (2) | 0 (1) | 82 (1) |
| 70+ years | 14 (1) | 118 (2) | 52 (6) | 19 (3) | 23 (6) | 2 (4) | 228 (3) |
| All ages | 69 (5) | 400 (8) | 107 (13) | 55 (8) | 49 (12) | 4 (9) | 685 (8) |
| At increased risk of severe COVID-19 illness |  |  |  |  |  |  |  |
| At least one condition | 305 (20) | 1450 (28) | 304 (38) | 186 (26) | 139 (34) | 13 (29) | 2398 (28) |
| Plus others older than: |  |  |  |  |  |  |  |
| 50 years | 362 (24) | 1937 (38) | 404 (50) | 247 (34) | 182 (44) | 18 (39) | 3150 (37) |
| 55 years | 339 (23) | 1762 (34) | 373 (46) | 225 (31) | 168 (41) | 16 (36) | 2884 (33) |
| 60 years | 323 (22) | 1637 (32) | 346 (43) | 209 (29) | 156 (38) | 15 (33) | 2686 (31) |
| 65 years | 329 (22) | 1634 (32) | 369 (46) | 213 (29) | 169 (41) | 16 (35) | 2729 (32) |
| 70 years | 319 (21) | 1557 (30) | 348 (43) | 203 (28) | 158 (39) | 15 (33) | 2599 (30) |

Figure 2. Percentage of global population at increased risk of severe COVID-19 disease (left panel) and \% change (right panel) when conditions are removed one at a time


The black shaded bar at the bottom represents the central estimate of the global population that are at increased risk. All other bars above show how this value changes when each of the conditions are removed one at a time. The most influential conditions are at the top of the bar chart and represent larger areas on the map shown on the right side.

Figure 3. Alternative version of Figure 3 in main paper showing the age-standardised proportion of population at increased risk and high risk of severe COVID-19 by country and region

Figure 3 in the main paper shows the share of the population at risk in different countries based on real-world differences in population structure and disease prevalence. This information is important when calculating the numbers that might need to be shielded or vaccinated but does not allow direct comparison of risks at equivalent ages in different countries. In this alternative version (see below), circles have been added to show the agestandardised share of the population at high risk (black circles) and increased risk (open circles). These assume each country has the same WHO standard reference population. 17 A low age-adjusted population at risk in countries with older populations (eg, Japan, Europe and Puerto Rico) helps to confirm that older age is the main reason why these countries have a high unadjusted population at risk. Similarly, a high age-adjusted population at risk in African countries with high HIV prevalence (eg, eSwatini, Lesotho) and small island nations with high diabetes prevalence (eg. Fiji, Mauritius) explains why these countries have a high unadjusted population at risk, despite having younger populations. Differences in demography can mask important differences in age-specific risks that may be revealed by age-standardisation. For example, in eSwatini and New Zealand the population at high risk is $5 \%$ in both countries, but when risks are compared for equivalent age groups (within the spreadsheet tool) the age-specific risks in eSwatini are more than double those in New Zealand (consistent with eSwatini having a higher age-adjusted population at high risk ie, $8 \%$ vs $3 \%$ ). Thus, although younger populations will tend to have a lower share of the population at risk than older populations, their risk at equivalent ages could still be higher.


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