

Original Paper

Cite this article: Herzog CM *et al* (2019). Pastoral production is associated with increased *peste des petits ruminants* seroprevalence in northern Tanzania across sheep, goats and cattle. *Epidemiology and Infection* **147**, e242, 1–9. <https://doi.org/10.1017/S0950268819001262>

Received: 12 February 2019

Revised: 10 May 2019

Accepted: 13 June 2019

Key words:

Epidemiology; peste-des-petits-ruminants; seroepidemiologic studies; Tanzania; animal husbandry

Author for correspondence:

Catherine M. Herzog,
E-mail: cqh5447@psu.edu

Pastoral production is associated with increased *peste des petits ruminants* seroprevalence in northern Tanzania across sheep, goats and cattle

C. M. Herzog¹, W. A. de Glanville², B. J. Willett³, T. J. Kibona⁴, I. M. Cattadori¹, V. Kapur¹, P. J. Hudson¹, J. Buza⁴, S. Cleaveland² and O. N. Bjørnstad¹

¹Pennsylvania State University, State College, Pennsylvania, USA; ²Institute of Biodiversity, Animal Health and Comparative Medicine, University of Glasgow, Glasgow, UK; ³MRC-University of Glasgow Centre for Virus Research, Glasgow, UK and ⁴Nelson Mandela African Institute of Science and Technology, Arusha, Tanzania

Abstract

Peste des petits ruminants virus (PPRV) causes a contagious disease of high morbidity and mortality in small ruminant populations globally. Using cross-sectional serosurvey data collected in 2016, our study investigated PPRV seroprevalence and risk factors among sheep, goats and cattle in 20 agropastoral (AP) and pastoral (P) villages in northern Tanzania. Overall observed seroprevalence was 21.1% (95% exact confidence interval (CI) 20.1–22.0) with 5.8% seroprevalence among agropastoral (95% CI 5.0–6.7) and 30.7% among pastoral villages (95% CI 29.3–32.0). Seropositivity varied significantly by management (production) system. Our study applied the catalytic framework to estimate the force of infection. The associated reproductive numbers (R_0) were estimated at 1.36 (95% CI 1.32–1.39), 1.40 (95% CI 1.37–1.44) and 1.13 (95% CI 1.11–1.14) for sheep, goats and cattle, respectively. For sheep and goats, these R_0 values are likely underestimates due to infection-associated mortality. Spatial heterogeneity in risk among pairs of species across 20 villages was significantly positively correlated (R^2 : 0.59–0.69), suggesting either cross-species transmission or common, external risk factors affecting all species. The non-negligible seroconversion in cattle may represent spillover or cattle-to-cattle transmission and must be investigated further to understand the role of cattle in PPRV transmission ahead of upcoming eradication efforts.

Introduction

Peste des petits ruminants virus (PPRV, family *Paramyxoviridae*, genus *Morbillivirus*) causes a contagious disease of high morbidity and mortality in sheep, goat and other small ruminant populations. Estimated mortality varies widely in the literature from <10% to as high as 90% in naïve herds [1, 2], but typically is reported between 50% and 80% [2]. The virus is spread by direct contact with infected hosts, aerosols, or fomites [2]; however, the infection can be prevented through vaccination [3]. First discovered in 1942 in Cote d'Ivoire [2], PPRV now has a broad geographic range and has been reported in over 70 countries in Africa, the Middle East and Asia. The Food and Agricultural Organization (FAO) estimates that PPRV threatens nearly 80% of the global small ruminant population and impacts the livelihood of over 330 million farmers who rely on small ruminant husbandry [4]. Low-income households can more easily afford small ruminants so the impact of PPRV on global opportunity for those in poverty is large. FAO estimates that PPRV causes approximately \$1.45–2.1 billion USD in global annual losses, with half of these losses impacting Africa and a quarter affecting Asia [4]. These economic losses are due to livestock mortality, impaired production and treatment and vaccination of morbid animals. Cost-benefit analyses estimate the net benefit of PPRV eradication at \$74.2 billion USD and a benefit-cost ratio of 33.8, based on a 15-year program [5].

Given its socio-economic importance and the availability of a vaccine that provides demonstrated protection up to 3 years [6], PPRV is an attractive target for eradication. In 2015, the FAO launched a global campaign to eradicate PPRV by 2030 [4, 7]. However, questions remain about specific ecological mechanisms driving PPRV transmission [8], including which risk factors and management systems are important for transmission and whether there are additional, unidentified domestic or wildlife reservoir hosts for PPRV. Given the historically high prevalence of the related rinderpest virus, PPRV has classically been thought of as a disease of small ruminants so cattle have not been routinely sampled during PPRV investigations under the presumption that they are not a competent host. However, there is mounting evidence that cattle can be infected and seroconvert [9–23], but to our knowledge, there have been no studies

© The Author(s) 2019. This is an Open Access article, distributed under the terms of the Creative Commons Attribution licence (<http://creativecommons.org/licenses/by/4.0/>), which permits unrestricted re-use, distribution, and reproduction in any medium, provided the original work is properly cited.

in which the exposure of sheep, goats and cattle have been investigated simultaneously to look for correlated epidemiologies. Furthermore, while there is increasing evidence of the importance of differing management (production) practices on PPRV transmission [24–26], these have not been explored in detail for PPRV among sheep, goats and cattle managed together.

To fill these gaps we conducted a cross-sectional serosurvey of sheep, goats and cattle managed together in the same households under differing production systems across 20 villages in northern Tanzania. In Tanzania, the first report of PPRV came from the northern border in late 2008. Within 2 years infection spread southward along known animal trade routes and PPRV is now present throughout the country [24]. Large-scale serosurveys conducted during 2008–2011 estimated PPRV seroprevalence to be 22%–45.8% in small ruminants in various regions of Tanzania [24, 25, 27, 28]. Government-led PPRV vaccination campaigns targeting sheep and goats were conducted in 2009–2011 in northern and southern districts and along livestock marketing routes and again in 2013 near Mikumi National Park, central Tanzania (Roos, 2016, unpublished MSc). However, no other major control efforts have occurred and PPRV is believed to be endemic across the majority of Tanzania.

Historically, cross-reactivity with rinderpest virus (RPV) has been a challenge for assessing PPRV seroprevalence during serosurveys, but this is no longer a significant concern as RPV was eliminated and vaccination ceased in Tanzania in 1997–1998 [29] and RPV was declared globally eradicated in 2011 [30]. Any small ruminant populations in Tanzania that may have been exposed to RPV infection have since been replaced (based on an average lifespan of approximately 1–3 years [24]) and although cattle are longer lived, only Tanzanian cattle 17 years or older in our study would have received a RPV vaccination. Thus, the current status and distribution of PPRV in Tanzania among sheep, goats and cattle can be clarified through serology. In this study, we address the following research aims: (i) to describe the current distribution of PPRV seroprevalence and the force of infection (*per capita* rate of infection) in northern Tanzania, (ii) to identify risk factors associated with PPRV seroprevalence and (iii) to determine if PPRV risk is correlated among sheep, goats and cattle.

Methods

Blood samples were collected from February to November 2016 from clinically healthy sheep, goats and cattle in 20 northern Tanzanian villages in Arusha and Manyara Regions as part of the Social, Economic and Environmental Drivers of Zoonoses (SEEDZ) study which originally sought to investigate brucellosis infection. A total of 7576 animals from 417 households were sampled and 404 household surveys were conducted. The sample size was based on the assumption of 40% within herd brucellosis prevalence and type 1 and 2 errors of 10%. Using multistage sampling, study villages were randomly selected from a spatially referenced list of 1012 villages (2012 national census, Tanzanian National Bureau of Statistics) using a generalised random tessellation stratified sampling (GRTS) approach. This provides a spatially balanced, probability-based sample [31, 32]. Prior to selection, all villages in the study area were classified according to dominant livelihood type by local experts (District Veterinary or Livestock Officers). ‘Pastoral’ (P) villages were those in which livestock rearing, rather than crop agriculture, was considered to be the primary livelihood activity while ‘agropastoral’ (AP) were those villages in which a mix of crop and livestock were

important. The GRTS procedure was stratified to select nine villages classified as agropastoral and 11 villages classified as pastoral (Fig. 1). Study villages comprised between two and four sub-villages, from which two or three were selected for inclusion in the study. The availability of animal restraint (livestock crush) was considered; otherwise, sub-village selection was random. Within each sub-village, a ‘central point’ sampling approach was adopted, in which livestock keepers were invited to bring their animals to a pre-selected point within the sub-village, such as a livestock crush or dip tank. The invitation included a 3-day advance notice to the sub-village chairperson and a meeting with the sub-village chairperson and as many compound leaders as possible 24 h before the event. These leaders notified livestock keepers of the event and encouraged as many herds as possible to attend and advised them to bring animals of all ages and both sexes. During the sampling event, a list of all attending households was generated and a maximum of 10 households was selected using a random number generator. A maximum of 10 cattle, 10 sheep and 10 goats was randomly selected for sampling from selected households to be able to detect herd level infection (estimated at 40% for the SEEDZ target pathogen, brucellosis). Animals were selected as every X^{th} animal (e.g. 2nd, 3rd, etc based on a number of animals present) moving through a livestock crush or, if corralled, by selecting at random points throughout the herd. Among cattle, five animals between 6 and 18 months of age and five animals over 18 months of age were selected. Among sheep and goats, five animals between 6 and 12 months of age and five animals over 12 months were selected. At least two of the animals were male in each of the groups and all animals were old enough to be expected to lack maternally derived PPRV immunity [33].

Animal age was assessed by dentition, namely incisor tooth eruption and wear (Fig. 2). This age estimation method was used because the exact recorded ages were not available. Maximum age cut-offs for sheep and goat age categories were: 1, 1.5, 2, 3, 5 and 5+ years [34, 35] and for cattle: 1.5, 2.5, 3.5, 4.5, 6 and 6+ years [36]. Among samples containing complete species, sex, age and location data, we tested a total of 7538 serum samples in duplicate using a commercial competitive ELISA kit (Pirbright Institute, Surrey, England) directed against the hemagglutinin protein of PPRV [37, 38]. Samples had been heat inactivated (56°C, 2 h) prior to shipment to the University of Glasgow for testing. We removed 42 samples (0.6% of total) from the analysis that were from households that self-reported PPRV vaccination in the past 24 months. However, we retained 427 samples (5.7% of total) that lacked a household survey from which to discern self-reported PPRV vaccination as the exclusion of these samples from analysis yielded qualitatively and quantitatively similar results. Our final analysis sample included 7496 animals, representing 2080 sheep, 2419 goats, 2997 cattle.

To explore the characteristics of PPRV transmission in greater detail, we calculated the basic reproductive ratio (R_0) and the force of infection (FOI). Since PPRV is fully immunizing (i.e. natural infection elicits an immune response detectable for the lifespan of the animal), we calculate R_0 according to $R_0 = 1/\text{seronegatives}$ for each species, management system and village under the assumptions of endemic equilibrium, random mixing and that seronegatives are fully susceptible [39]. We calculated the FOI, the rate of infection of susceptible hosts, using the catalytic framework [40]. The catalytic model provides the framework for calculating the FOI from age-specific seroprevalence data from cross-sectional surveys [40–42]. Given that the rate at which immunity builds up with age depends on rates of circulation,

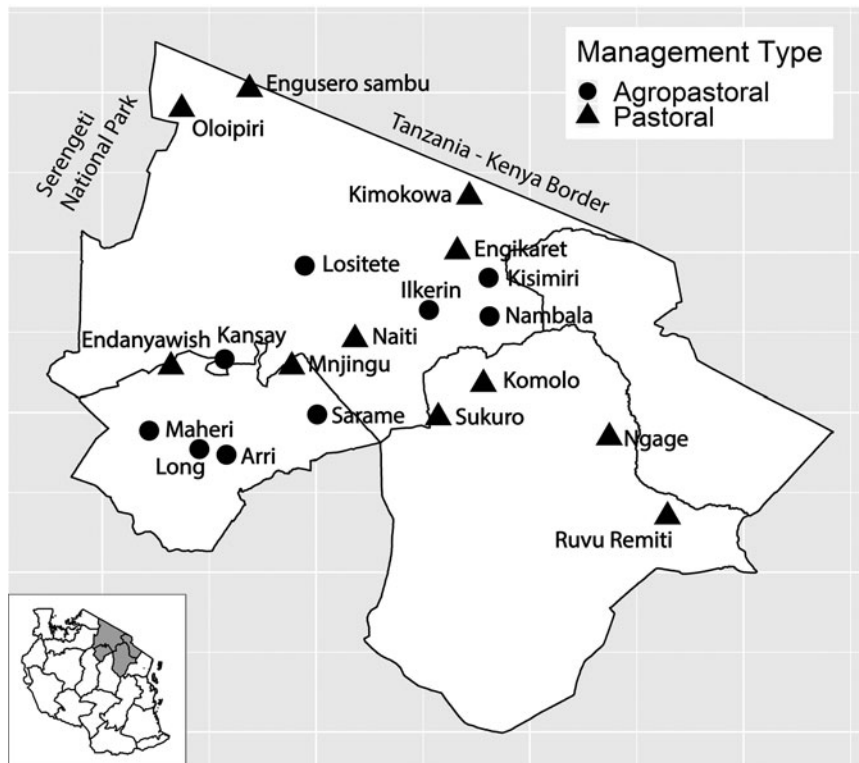


Fig. 1. Study villages in northern Tanzania and village management system type. Inset of Tanzania highlights the three regions in main figure.

this approach allows us both to elucidate the rate of circulation (FOI, λ) and to determine which age classes are most important for continued transmission. According to the catalytic model, age-specific seroprevalence, $P(a)$, can be expressed as:

$$\frac{dP(a)}{da} = \lambda(a)(1 - P(a)) \quad (1)$$

where $(1 - P(a))$ is the proportion of susceptible hosts of age a and $\lambda(a)$ is the age-specific FOI. Integrating and rearranging yields

$$P(a) = 1 - \exp\left[-\int_0^a \lambda(a) da\right] \quad (2)$$

which describes how the cumulative FOI up to a given age will deplete a susceptible cohort, where $P(a)$ represents the probability of having been infected before age a . Age-specific seroprevalence curves are direct empirical observations on this probability.

For this study, we are primarily interested in contrasting PPRV circulation among species and between management systems; therefore, we make the simplifying assumption of a constant FOI across ages. This assumption averages across age-specific heterogeneities and allows us to simultaneously estimate the FOI and explore the significance of risk factors on PPRV circulation. With a constant FOI, equation (2) simplifies to:

$$P(a) = 1 - \exp[-\lambda a] \quad (3)$$

which upon rearrangement corresponds to the expectation of a binomial generalised linear model with a complementary log-log link:

$$\log(-\log(1 - P(a))) = \log(\phi) + \log(a) \quad (4)$$

In (4), the intercept, $\log(\phi)$, represents an estimate of the log FOI as long as we use $\log(a)$ as a regression offset (i.e. has a regression coefficient fixed at unity). We applied this approach to a combined, all-species model with species as a fixed effect as well as to each of three individual-species models. We used sex and management system type as additional covariates. We used village as a random effect to explore spatial covariation in the PPRV risk among species. The effect of interactions (species \times management type, sex \times management type, sex \times species) were also explored (Table S1). We identified the best fitting all-species model and individual-species models using the Akaike Information Criterion [43]. For the all-species model, the reference group was a male cow in an agropastoral system. We calculated the village-level random effects from individual-species generalised linear mixed models using the lme4 package. Spatial variation in FOI rates and correlations in FOI between pairs of species were investigated via the random effects. Spatial autocorrelation in FOI was tested for using Moran's I statistic. All analyses were conducted in R version 3.5.3 [44].

All participants provided written informed consent or assent. Questionnaires collected household demographic and socio-economic information, assessed knowledge about common practices for livestock management and disease and documented animal movement in the region. In this study, the questionnaires were used solely to determine the PPRV vaccination status of household herds in the past 12 months. The protocols, questionnaire and consent/assent procedures were approved by the following ethical review committees: Kilimanjaro Christian Medical Centre (KCMC/832) and National Institute of Medical Research (NIMR/2028) in Tanzania; College of Medical, Veterinary and Life Sciences, University of Glasgow in the UK. Approval for animal work was provided by the Clinical Research Ethics Committee at the University of Glasgow School of Veterinary Medicine (39a/15).

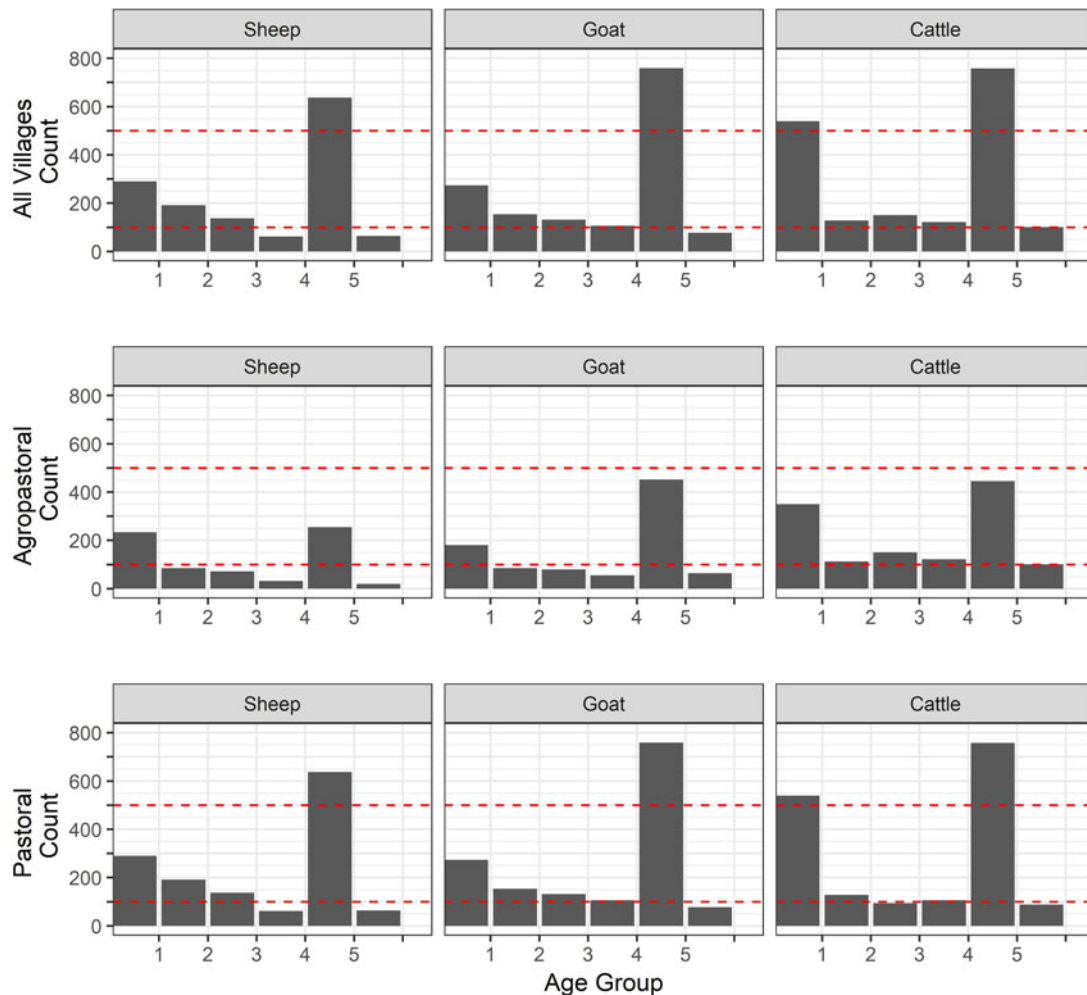


Fig. 2. Age distribution of sheep, goats and cattle. Age was measured in the animals by dentition (age group 1 = temporary teeth, 2 = one pair, 3 = two pair, 4 = three pair, 5 = 4 pair, 6 = full and worn (over 5 years old)). Dashed lines are included for comparison and all set at 100 and 500.

Results

Demographic characteristics of the 7496 samples analysed are presented in Table 1. In this section, ranges indicate 95% exact confidence intervals. The overall observed PPRV seroprevalence was 21.1% (20.1%–22.0%). Seroprevalence was 26.2% in sheep (24.3%–28.1%), 28.8% in goats (27.0%–30.6%) and 11.3% in cattle (10.2%–12.5%). There was significant variation within specific management systems: 0.8%–20.6% in agropastoral (AP) systems (overall AP: 5.8%, 5.0%–6.7%) and 3.3%–47.2% in pastoral (P) villages (overall P: 30.7%, 29.3%–32.0%) (Fig. 3, Fig. S1). Additionally, within each species, seroprevalence varied significantly by management system type (sheep AP: 6.0%, $P = 36.4\%$; goats AP: 7.0%, $P = 42.1\%$; cattle AP: 4.9%, $P = 16.1\%$).

Based on seroprevalence, reproductive number estimates were 1.36 (1.32–1.39), 1.40 (1.37–1.44) and 1.13 (1.11–1.14) for sheep, goats and cattle, respectively. Stratified by management system, we see higher R_0 values for each species in the pastoral management systems: sheep AP: 1.06 (1.05–1.09), $P = 1.57$ (1.51–1.64); goats AP: 1.08 (1.06–1.10), $P = 1.73$ (1.65–1.81); cattle AP: 1.05 (1.04–1.07), $P = 1.19$ (1.17–1.22). At a village level, R_0 estimates range from 1.00 to 2.77 for sheep, 1.00–3.18 for goats and 1.00–1.53 for cattle, with 1, 4 and 2 villages (out of 20) having a result of 1.00 in each species, respectively. These estimates are at or above the threshold

for invasion but are likely to be underestimated as they have not been corrected for infection-associated mortality and PPR is known to cause high mortality in small ruminants (but not in cattle). Although our study lacked case fatality data (sampled animals were clinically healthy), one can calculate plausible biases introduced by infection-associated mortality. For example, for a 50% case mortality rate, one would expect R_0 to be underestimated by 25%–45% depending on the true value of R_0 (between 2 and 10) using the $1/\text{seronegatives}$ calculation. The bias results because infection killed animals are not present among the seropositives.

For the best fitting all-species model containing one interaction (species:management, see Table S1), sex, management and the species:management interaction were significant, with the change in management system exhibiting the strongest hazard ratio (HR: 3.8, 1.4–10.3) (Table 2). The species:management interaction is implicit in the individual-species models that the rest of our results are based on. In individual-species models, the hazard ratio of seroconversion between pastoral and agropastoral systems was 9.4 for sheep (3.6–24.55), 9.5 for goats (3.1–29.4) and 4.0 for cattle (1.6–10.1) (Table 2).

Spatial variation in FOI is presented in Figure 4. There was significant spatial autocorrelation in village-level FOI for goats and cattle but not for sheep. Spatial FOI were strongly correlated

Table 1. Sample population characteristics

	<i>n</i> (%)	PPRV Seropositive <i>n</i> (%; 95% CI)
Total	7496	1580 (21.1, 20.2–22.0)
Species		
Sheep	2080 (27.7)	545 (26.2, 24.3–28.1)
Goats	2419 (32.3)	696 (28.8, 27.0–30.6)
Cattle	2997 (40)	339 (11.3, 10.2–12.5)
Sex		
Female	5508 (73.5)	1393 (25.3, 24.1–26.4)
Male	1988 (26.5)	187 (9.4)
Management system		
Agropastoral (AP)	2898 (38.7)	169 (5.8, 5.0–6.7)
Pastoral (P)	4598 (61.3)	1411 (30.7, 29.4–32.0)
Age Group ^a		
Temporary teeth	1869 (24.9)	106 (5.7, 4.7–6.8)
1 pair	757 (10.1)	68 (9.0, 7.0–11.2)
2 pair	666 (8.9)	88 (13.2, 10.7–16.0)
3 pair	483 (6.4)	97 (20.1, 16.6–23.9)
4 pair	3307 (44.1)	1088 (32.9, 31.3–34.5)
Full mouth + worn	414 (5.5)	133 (32.1, 27.6–36.9)
Breed		
Indigenous breed	7122 (95.0)	1535 (21.6, 20.6–22.5)
Cross breed	370 (4.9)	44 (11.9, 8.8–15.6)
Exotic breed	4 (0.05)	1 (25.0, 0.6–80.6)
Village ^b		
Arri (AP)	125 (4.3)	2 (1.6, 0.2–5.6)
Ilkerin (AP)	343 (11.8)	45 (13.1, 9.7–17.2)
Kansay (AP)	458 (15.8)	6 (1.3, 0.5–2.8)
Kisimiri (AP)	228 (7.9)	47 (20.6, 15.6–26.5)
Long (AP)	261 (9.0)	2 (0.8, 0.09–2.7)
Lositete (AP)	260 (9.0)	32 (12.3, 8.6–16.9)
Maheri (AP)	511 (17.6)	1 (0.2, 0.00005–1.0)
Nambala (AP)	352 (12.2)	22 (6.3, 4.0–9.3)
Sarame (AP)	360 (12.4)	12 (3.3, 1.7–5.8)
Endanyawish (P)	213 (4.6)	7 (3.3, 1.3–6.7)
Engikaret (P)	529 (11.5)	151 (28.5, 24.7–32.6)
Engusero Sambu (P)	351 (7.6)	55 (15.7, 12.0–20.0)
Kimokowa (P)	443 (9.6)	153 (34.5, 30.1–39.1)
Komolo (P)	413 (9.0)	102 (24.7, 20.6–29.2)

(Continued)

Table 1. (Continued.)

	<i>n</i> (%)	PPRV Seropositive <i>n</i> (%; 95% CI)
Mnjingu (P)	350 (7.6)	78 (22.3, 18.0–27.0)
Naiti (P)	584 (12.7)	271 (46.4, 42.3–50.5)
Ngage (P)	356 (7.7)	152 (42.7, 37.5–48.1)
Oloipiri (P)	401 (8.7)	130 (32.4, 27.9–37.2)
Ruvu Remitii (P)	525 (11.4)	248 (47.2, 42.9–51.6)
Sukuro (P)	433 (9.4)	64 (14.8, 11.6–18.5)

^aAge groups by dentition correspond to approximately the following ages in cattle: temporary teeth: 1 month to 1.5 years; 1 pair: 1.5–2.5 years; 2 pair: 2.5–3.5 years, 3 pair: 3.5–4.5 years, 4 pair: 4.5–7 years; full + worn: above 7 years. For sheep and goats: temporary teeth: <1 year; 1 pair: 1–1.5 years; 2 pair: 1.5–2 years, 3 pair: 2–3 years, 4 pair: 3–5 years; full + worn: above 5 years.

^bDenominators for the villages were the *n* of the corresponding management system type. CI, confidence interval.

among species (Fig. 5) after adjustment for sex and management system and showed a significant positive association ($P < 0.001$) between each pair of species: goats and sheep (R^2 : 0.71), sheep and cattle (R^2 : 0.68) and goats and cattle (R^2 : 0.61). This pattern is consistent with similar within-village transmission rates between each species and the presence of either cross-species transmission and/or common, external risk factors affecting all species.

Discussion

Our study has (i) described the current distribution of PPRV seroprevalence and force of infection (FOI) among sheep, goats and cattle in northern Tanzania, (ii) identified management system and sex (overall in sheep and goats) as significant risk factors for PPRV seroprevalence and (iii) provided evidence of transmission patterns among sheep, goats and cattle consistent with cross-species circulation. PPRV seropositivity was found in all species but was higher in small ruminants than in cattle, higher in females than males and increased across dentition-based age groups as would be expected with a fully immunizing infection and consistent with an endemic pattern of circulation. PPRV exposure was widespread and varied across the northern Tanzanian landscape and seroprevalence was significantly higher for all three species under pastoral management than under agropastoral management. Significant spatial autocorrelation in FOI after controlling for sex and management system was seen in cattle and goats, but not in sheep, suggesting other landscape effects (market proximity and attendance, veterinary service access) may impact PPRV seroprevalence. Lastly, our study found a significant positive association in the spatial variation in FOI between each pair of species suggesting cross-species transmission is occurring in sheep, goats and cattle and/or that common external forces are affecting transmission within each species. Given the multi-host aetiology seen with other morbilliviruses, such as canine distemper [45] and phocine distemper viruses [46], the former explanation is an interesting hypothesis for future study.

Our finding that management system type has a significant effect on PPRV seroprevalence parallels previous findings, as does our find that sex (in sheep and goats) has a significant effect (Supplementary Text). In Tanzania, Kivaria *et al.* [24] and Swai *et al.* [25] reported a significant effect of pastoral management systems (odds ratios: 3.56 (95% CI 3.02–4.21) and 2.08 (95% CI

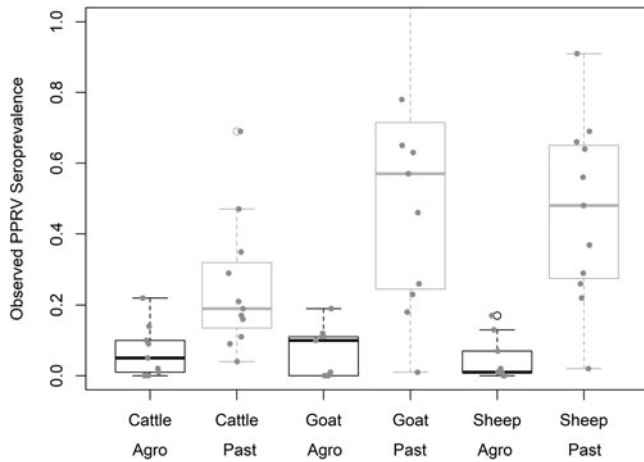


Fig. 3. Variation in village PPRV seroprevalence by species and management type. Agro = Agropastoral, Past = Pastoral.

1.55–2.78), respectively). Beyond Tanzania, pastoral management systems have also been associated with an increased risk of PPRV transmission in Ethiopia and Sudan ([26], Dejene 2016 unpublished MSc). Density dependence in PPRV transmission remains to be investigated and herd size may be a contributor to the difference in PPRV infection hazard between management system types as pastoral herds are generally larger than agropastoral herds. Additionally, management type variation should be investigated at finer geographic scales. If household-level management classification were available it may be possible to correlate household specific management practices obtained by questionnaire to variation in PPRV transmission risk within the classification. Additionally, pastoral livelihoods are changing, with crop cultivation, migration and wage jobs on the rise among the Masaai in northern Tanzania [47]; thus, getting a finer view of household practices is essential. Variation in herd size, age structure, mixing patterns (grazing, confinement, watering, migration practices), acquisition rates and sources of new stock, wildlife exposure and access to veterinary services of herds in each management system should be explored. Opportunities exist for increasing veterinary service provisions that improve access to routine, preventative veterinary care including education (i.e. promotion of changing common production system practices in ways that reduce transmission risk) and vaccination.

Our work adds to a growing literature that has investigated PPRV in cattle. In our investigation, cattle were managed in the same households as the sheep and goats sampled. Our study found an overall cattle seroprevalence of 11.3% (AP: 4.9%, P: 16.1%) which is comparable to the seroprevalence range reported by most other studies that have tested cattle (range: 1.8–18%, $n = 24$ –1158) [9–20, 23] (Kundu 2013, unpublished MSc), save two which reported cattle seroprevalence at 41.9% ($n = 43$, by our calculations a 95% CI 27.0–57.9) [10] and 42% ($n = 1,000$, by our calculations a 95% CI 38.9–45.1) [15]. Additionally, experimental infections [21, 22] have demonstrated that cattle can be infected by PPRV via inoculation or contact with infected small ruminants. Cattle are thought to be a less competent host for PPRV; however, it is not known if cattle are able to transmit PPRV or whether infections are merely spillovers from small ruminants. This is an area where carefully controlled experimental studies are needed. Prior to RPV eradication, it is possible that PPRV was blocked in cattle due to cross-immunity between the viruses

Table 2. Generalised linear mixed model results for all-species and individual-species models

Fixed Effect	All Species Model ^a			Sheep Model ^a			Goat Model ^a			Cattle Model ^a		
	HR	95% CI ^b	P-Value	HR	95% CI ^b	P-Value	HR	95% CI ^b	P-Value	HR	95% CI ^b	P-Value
Intercept	0.008	(0.004, 0.02)	<0.001	0.008	(0.003, 0.02)	<0.001	0.009	(0.003, 0.02)	<0.001	0.008	(0.004, 0.02)	<0.001
Sex	1.5	(1.2, 1.7)	<0.001	1.6	(1.2, 2.2)	<0.001	1.4	(1.1, 1.8)	0.004	1.3	(1.0, 1.8)	0.09
Goat	1.2	(0.8, 1.7)	0.31	-	-	-	-	-	-	-	-	-
Sheep	0.9	(0.6, 1.4)	0.78	-	-	-	-	-	-	-	-	-
Pastoral Management	3.8	(1.4, 10.3)	0.008	9.4	(3.6, 24.4)	<0.001	9.5	(3.1, 29.4)	<0.001	4.0	(1.6, 10.1)	0.004
Goat : Management	2.5	(1.7, 3.6)	<0.001	-	-	-	-	-	-	-	-	-
Sheep : Management	2.9	(1.9, 4.5)	<0.001	-	-	-	-	-	-	-	-	-

HR, hazard ratio estimate.

^aReference group: male cattle in an agropastoral system. See Table S1 for interaction (species:management) selection.

^bWald confidence interval.

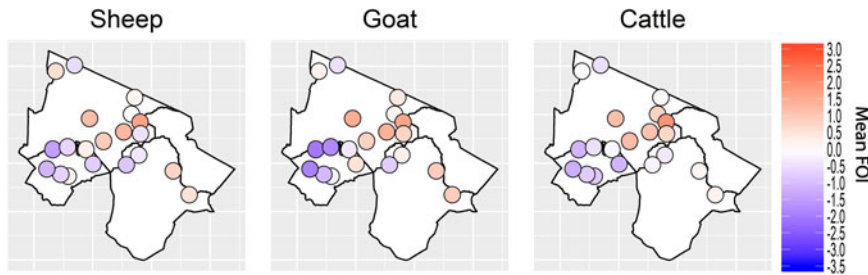


Fig. 4. Spatial variation in village-level standardised force of infection (FOI) estimates. Blue colouring indicates a village FOI that was lower than the mean FOI (standardised to zero) and red colouring indicates a village FOI that was higher than the standardised mean FOI. There was strong evidence of spatial autocorrelation (Moran's I) in FOI for goats and cattle ($P < 0.001$) but weak evidence in sheep ($P = 0.08$).

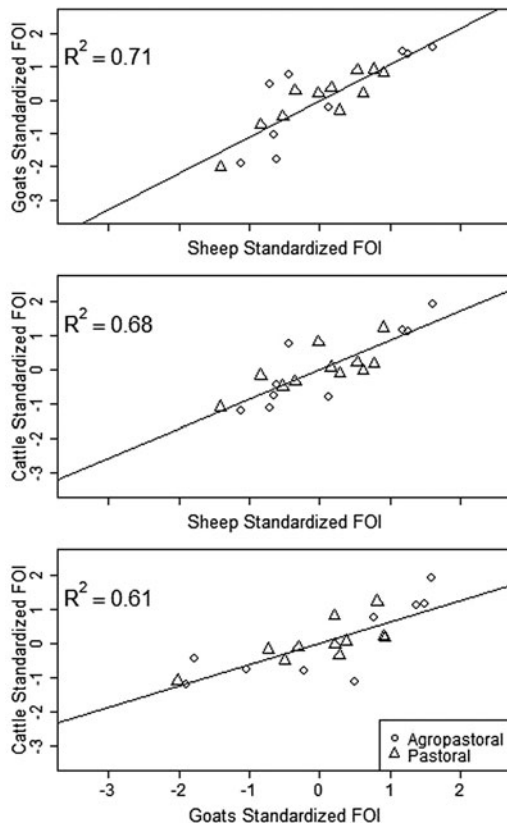


Fig. 5. Village-level standardised force of infection (FOI) estimates by species pairs. Mean FOI is standardised to zero and is in middle of both axes. Where the village FOI was higher or lower than the mean FOI for one species, it was also higher or lower than the mean for the second species plotted. All correlations are significant ($P < 0.001$).

[48–51]. Post RPV eradication, PPRV may have further adapted to cattle. It is possible that PPRV spillover from cattle to small ruminants may be an additional route of transmission that is currently present or may develop as PPRV adapts to the vacated niche left by RPV [8, 52]. Our analysis of village-level FOI adds support to several alternative hypotheses regarding cattle transmission: (i) that cross-species PPRV transmission is occurring among sheep, goats and cattle, (ii) that independent, within-species chains of transmission are enhanced or reduced by common external factor(s) affecting all species, or (iii) both are occurring. Additionally, to resolve the role of cattle in PPRV transmission (and the directionality of transmission), longitudinal field investigations of viral excretion, strain characterisation by species and species contact patterns by management system should be pursued in multi-species herds.

Our study methods assumed PPRV endemicity in Tanzania. Our results showed a pattern of increasing seroprevalence with age, which is consistent with an endemic infection pattern in that individuals' cumulative exposure increases with age. Given that PPRV has arrived in Tanzania relatively recently, future studies with increased temporal and geographic granularity may be able to confirm PPRV endemicity in Tanzania or resolve areas where PPRV is still epidemic. Cross-sectional analyses such as ours cannot distinguish stable endemicity from recurrent epidemics, but statistical investigations into the impact of epidemics on FOI and R_0 estimates for mumps and rubella derived from serological data have demonstrated that they remain mostly insensitive to regular epidemic cycles [53]. Furthermore, the authors expected the same conclusions for irregular epidemics with a period less than the average age of infection. We expect these results to be generalisable to PPRV; hence, we believe that our assumption of endemicity is a reasonable first approximation in this analysis.

Our model purposely assumed an age-invariant FOI to enable us to simultaneously estimate the FOI and explore the significance of risk factors on PPRV circulation in different host species. However, FOI may vary by age. Using an age-varying FOI model (e.g. piecewise-constant model) could identify which age cohorts experience the highest burden of PPRV circulation and is the subject of our future research. The results from an age-varying model could also better inform hypotheses about the mechanisms driving the significant difference that our study and several other studies (Supplementary Text) saw by sex: namely, females had significantly higher seroprevalence than males. For example, an age-varying model may reveal that females experience higher FOI during peak reproductive ages. If so, this could be due to biological factors, such as increased susceptibility due to immune changes during pregnancy, or social factors, such as mixing more with young animals under a year old who, after losing maternal immunity and prior to vaccination, become infected and pass PPRV to their mothers. Additionally, infection-associated mortality, seasonal susceptible influxes and preferential animal slaughter by a specific species, sex, age, or weight (age proxy) as part of cultural or religious festivals may impact dynamic patterns and be revealed by this model.


Our study did not collect animal contact data and assumed homogenous mixing among animals. This is likely a plausible first approximation as herds are often grazed and watered together during the day and although adult cattle are often separated from sheep and goats at night, calves are not. Study animals were kept in the areas they were sampled in but may have originated elsewhere. Detailed studies of contact patterns and trade movements in various production systems are needed to specify alternate mixing patterns and improve epidemiological parameter estimates. Our study sampling methodology strived to reduce

selection bias by using random selection at all stages of sampling and stressing that all households in the area should bring all animals for sampling. However, there was evidence that households closer to the sampling point, with larger herd sizes and lower socioeconomic status were more likely to attend (unpublished observation). There may have also been animal selection based on health (e.g. sick or frail animals purposely brought/left at home). Lastly, our work assumes that a positive cELISA result indicates past PPRV infection and current protection and a negative result indicates no past exposure and current susceptibility. Misclassifications would result in underestimation of seroprevalence and our R_0 estimates. The BDSL cELISA test in our study has been previously used on Tanzanian cattle and wildlife samples [16]. With a threshold of 50%, the test demonstrated high specificity; however, sensitivity could be improved for all species [38], especially cattle (possibly by lowering the threshold <50% [54]) and other unusual PPRV hosts such as camels [9] and wildlife. As the test may be less sensitive in cattle, our seroprevalence results are likely to be underestimated.

Our study has highlighted the importance of the pastoral production system and the need to further investigate the role of cattle in PPRV transmission risk. Resolving the ecological mechanisms driving PPRV transmission will lead to improved prevention strategies and better inform upcoming eradication campaigns.

Supplementary material. The supplementary material for this article can be found at <https://doi.org/10.1017/S0950268819001262>

The anonymised dataset used in this study will be made available on <http://researchdata.gla.ac.uk/>

Author ORCIDs.  C. M. Herzog, 0000-0003-3021-888X.

Acknowledgements. We thank all members of the SEEDZ field and laboratory team, including Kunda Mnzava, Eurphrasia Mariki, Zanuni Kweka, Rigobert Tarimo, Fadhili Mshana, Mamus Toima, Tauta Maapi, Raymond Mollé, Matayo Lazaro, Ephrasia Hugo, Nelson Amana, Elizabeth Kasagama and Victor Moshi. C. M. H. was supported by the Bill & Melinda Gates Foundation grant 'Programme For Enhancing the Health and Productivity of Livestock (PEHPL)' (OPP1083453). Additional funding for this study was provided by the Biotechnology and Biological Sciences Research Council, the Department for International Development, the Economic & Social Research Council, the Medical Research Council, the Natural Environment Research Council and the Defense Science & Technology Laboratory, under the Zoonoses and Emerging Livestock Systems (ZELS) programme (BB/L018926/1).

Conflict of interest. None.

References

1. **Global Strategy for the Control and Eradication of PPR** (2015) World Organization for Animal Health & Food and Agricultural Organization of the United Nations. <http://www.fao.org/3/a-i4460e.pdf>
2. **Munir M** ed (2015) *Peste des Petits Ruminants Virus*. Heidelberg: Springer.
3. **Kumar N et al.** (2017) Advances in peste des petits ruminants vaccines. *Veterinary Microbiology* **206**, 91–101.
4. **Global control and eradication of peste des petits ruminants** (2015) World Organization for Animal Health & Food and Agricultural Organization of the United Nations. <http://www.fao.org/3/a-i4477e.pdf>
5. **Jones BA et al.** (2016) The economic impact of eradicating peste des petits ruminants: a benefit-cost analysis. *PLoS ONE* **11**, 1–18.
6. **Zahur AB et al.** (2014) Peste des petits ruminants vaccine (Nigerian strain 75 / 1) confers protection for at least 3 years in sheep and goats. *Journal of Biosciences and Medicines* **2**, 27–33.
7. **Peste des Petits Ruminants (PPR) A challenge for small ruminant production** (2009) Food and Agricultural Organization of the United Nations. http://www.fao.org/ag/againfo/resources/documents/AH/PPR_flyer.pdf
8. **Baron MD et al.** (2017) Future research to underpin successful peste des petits ruminants virus (PPRV) eradication. *Journal of General Virology* **98**, 2635–2644.
9. **Abraham G et al.** (2005) Antibody seroprevalences against peste des petits ruminants (PPR) virus in camels, cattle, goats and sheep in Ethiopia. *Preventive Veterinary Medicine* **70**, 51–57.
10. **Khan HA et al.** (2008) The detection of antibody against peste des petits ruminants virus in sheep, goats, cattle and buffaloes. *Tropical Animal Health and Production* **40**, 521–527.
11. **Ngangnou A et al.** (1996) Evaluation de la protection vaccinale contre la peste bovine au cameroun III, evaluation globale. *Revue d'élevage et de Médecine Vétérinaire des Pays Tropicaux* **49**, 18–22.
12. **Toukara K et al.** (1996) Epidemiologie de la peste des petits ruminants (PPR) et de la peste bovine au Mali: enquetes serologiques. *Revue d'élevage et de Médecine Vétérinaire des Pays Tropicaux* **49**, 273–277.
13. **Balamurugan V et al.** (2014) Prevalence of peste des petits ruminants in goats in North-East India. *Virus Disease* **25**, 488–492.
14. **Intisar KS et al.** (2017) Peste des petits ruminants infection in domestic ruminants in Sudan. *Tropical Animal Health and Production* **49**, 747–754.
15. **Ali WH** (2018) Serological investigations of peste des petits ruminants among cattle in the Sudan. *Tropical Animal Health and Production* **51** (3), 655–659.
16. **Lembo T et al.** (2013) Peste des petits ruminants infection among cattle and wildlife in northern Tanzania. *Emerging Infectious Diseases* **19**, 2037–2040.
17. **Rashid A, Asim M and Hussain A** (2008) Seroprevalence of peste des petits ruminants (PPR) virus in goats, sheep, cattle, at Livestock Production Research Institute Bahadurnagar Okara. *Journal of Animal and Plant Sciences* **18**, 114–116.
18. **Balamurugan V et al.** (2012) Seroprevalence of peste des petits ruminants in cattle and buffaloes from southern peninsular India. *Tropical Animal Health and Production* **44**, 301–306.
19. **Albayrak H and Gür S** (2010) A serologic investigation for peste des petits ruminants infection in sheep, cattle and camels (*Camelus dromedarius*) in Aydin province, West Anatolia. *Tropical Animal Health and Production* **42**, 151–153.
20. **Haroun M et al.** (2002) Detection of antibodies against peste des petits ruminants virus in sera of cattle, camels, sheep and goats in Sudan. *Veterinary Research Communications* **26**, 537–541.
21. **Sen A et al.** (2014) Detection of subclinical peste des petits ruminants virus infection in experimental cattle. *Virus Disease* **25**, 408–411.
22. **Dardiri A, DeBoer C and Hamdy F** (1976) *Response of American goats and cattle to peste des petits ruminants virus. Proceedings of the 19th Annual Meeting of the American Association of Veterinary Laboratory Diagnosticians*. Miami Beach, Florida: The American Association of Veterinary Laboratory Diagnosticians, p. 337–344.
23. **Abubakar M et al.** (2017) Serological detection of antibodies to peste des petits ruminants virus in large ruminants. *Transboundary and Emerging Diseases* **64**, 513–519.
24. **Kivaria FM et al.** (2013) The incursion, persistence and spread of peste des petits ruminants in Tanzania: epidemiological patterns and predictions. *The Onderstepoort Journal of Veterinary Research* **80**, 1–10.
25. **Swai ES et al.** (2009) Prevalence and distribution of peste des petits ruminants virus antibodies in various districts of Tanzania. *Veterinary Research Communications* **33**, 927–936.
26. **Salih HAM et al.** (2014) Seroprevalence and risk factors of peste des petits ruminants in sheep and goats in Sudan. *Journal of Advanced Veterinary and Animal Research* **1**, 42–49.
27. **Muse EA et al.** (2012) Epidemiological investigation into the introduction and factors for spread of peste des petits ruminants, southern Tanzania. *Onderstepoort Journal of Veterinary Research* **79**, 2–7.
28. **Mbyuzi AO et al.** (2014) Sero-prevalence and associated risk factors of peste des petits ruminants and contagious caprine pleuro-pneumonia in

- goats and sheep in the southern zone of Tanzania. *Preventive Veterinary Medicine* **116**, 138–144.
29. **Inter-African Bureau for Animal Resources** (2011) *The Eradication of Rinderpest from Africa: A Great Milestone*. Nairobi, Kenya: Inter-African Bureau for Animal Resources.
 30. **The Global Rinderpest Eradication Programme** (2011) Progress report on rinderpest eradication: Success stories and actions leading to the June 2011 Global Declaration. http://www.fao.org/ag/againfo/resources/documents/AH/GREP_flyer.pdf
 31. **Stevens DL and Olsen AR** (2003) Variance estimation for spatially balanced samples of environmental resources. *Environmetrics* **14**, 593–610.
 32. **Stevens DL and Olsen AR** (2004) Spatially balanced sampling of natural resources. *Journal of the American Statistical Association* **99**, 262–278.
 33. **Bodjo C et al.** (2006) Assessment of the duration of maternal antibodies specific to the homologous peste des petits ruminants vaccine Nigeria 75/1 in Djallonke lambs. *Biokemistri* **18**, 99–103.
 34. **Mitchell T** (2003) AgFact A7.2.2 How to tell the age of goats. *Division of Animal Production, Dubbo, NSW Agriculture*. doi: https://www.dpi.nsw.gov.au/_data/assets/pdf_file/0007/178477/age-of-goats.pdf.
 35. **Casburn G** (2016) PrimeFact: How to tell the age of sheep. *Department of Primary Industries, NSW Agriculture*. doi: https://www.dpi.nsw.gov.au/_data/assets/pdf_file/0004/179797/aging-sheep.pdf.
 36. **Turton JA** (1999) How to estimate the age of cattle. Department of Agriculture, Forestry, and Fisheries, Republic of South Africa. <https://www.daff.gov.za/daffweb3/Portals/0/InfoPaks/cattleEstimateAge.pdf>
 37. **Anderson J, McKay JA and Butcher RN** (1991) The use of monoclonal antibodies in competitive ELISA for the detection of antibodies to rinderpest and peste des petits ruminants viruses. *The Sero-Monitoring of Rinderpest throughout Africa: Phase I*, 45–53. doi: https://inis.iaea.org/Collection/NCLCollectionStore/_Public/22/084/22084087.pdf?r=1&r=1.
 38. **Couacy-Hymann E et al.** (2007) Comparison of two competitive ELISAs for the detection of specific peste-des-petits-ruminant antibodies in sheep and cattle populations. *African Journal of Biotechnology* **6**, 732–736.
 39. **Anderson R. M. and May RM** (1992) *Infectious Diseases of Humans: Dynamics and Control*. New York: Oxford University Press.
 40. **Hens N et al.** (2012) *Modeling Infectious Disease Parameters Based on Serological and Social Contact Data: A Modern Statistical Perspective*. New York: Springer.
 41. **Hens AN et al.** (2010) Seventy-five years of estimating the force of infection from current status data. *Epidemiology and Infection* **138**, 802–812.
 42. **Muench H** (1934) Derivation of rates from summation data by the catalytic curve. *Journal of the American Statistical Association* **29**, 25–38.
 43. **Akaike H** (1974) A new look at the statistical model identification. *IEEE Transactions on Automatic Control* **19**, 716–723.
 44. **R Core Team** (2018). R: A language and environment for statistical computing. R Foundation for Statistical Computing, Vienna, Austria. URL <https://www.R-project.org/>.
 45. **Craft ME et al.** (2008) Dynamics of a multihost pathogen in a carnivore community. *Journal of Animal Ecology* **77**, 1257–1264.
 46. **Jo WK, Osterhaus AD and Ludlow M** (2018) Transmission of morbilliviruses within and among marine mammal species. *Current Opinion in Virology* **28**, 133–141.
 47. **McCabe JT et al.** (2014) Livelihood diversification through migration among a pastoral people: contrasting case studies of maasai in northern Tanzania. *Human Organization* **73**, 389–400.
 48. **Couacy-Hymann E, Bodjo SC and Danho T** (2006) Interference in the vaccination of cattle against rinderpest virus by antibodies against peste des petits ruminants (PPR) virus. *Vaccine* **24**, 5679–5683.
 49. **Mariner JC et al.** (1993) The use of thermostable vero cell-adapted rinderpest vaccine as a heterologous vaccine against peste des petits ruminants. *Research in Veterinary Science* **54**, 212–216.
 50. **Taylor W** (1979) Protection of goats against peste-des-petits ruminants with attenuated rinderpest virus. *Research in Veterinary Science* **27**, 321–324.
 51. **Couacy-Hymann E et al.** (1995) Protection of goats against rinderpest by vaccination with attenuated peste des petits ruminants virus. *Research in Veterinary Science* **59**, 106–109.
 52. **Lloyd-Smith JO** (2013) Vacated niches, competitive release and the community ecology of pathogen eradication. *Philosophical Transactions of the Royal Society of London. Series B, Biological Sciences* **368**, 1–12.
 53. **Whitaker HJ and Farrington CP** (2004) Estimation of infectious disease parameters from serological survey data: the impact of regular epidemics. *Statistics in Medicine* **23**, 2429–2443.
 54. **Singh RP et al.** (2004) Development of a monoclonal antibody based competitive-ELISA for detection and titration of antibodies to peste des petits ruminants (PPR) virus. *Veterinary Microbiology* **98**, 3–15.