Antenatal blood pressure for prediction of pre-eclampsia, preterm birth, and small for gestational age babies: development and validation in two general population cohorts

Corrie Macdonald-Wallis,1,2 Richard J Silverwood,3,6 Bianca L de Stavola,3,6 Hazel Inskip,5,6 Cyrus Cooper,5,6,7 Keith M Godfrey,5,6 Sarah Crozier,5 Abigail Fraser,1,2 Scott M Nelson,8 Debbie A Lawlor,1,2 Kate Tilling1,2

ABSTRACT

STUDY QUESTION
Can routine antenatal blood pressure measurements between 20 and 36 weeks' gestation contribute to the prediction of pre-eclampsia and its associated adverse outcomes?

METHODS
This study used repeated antenatal measurements of blood pressure from 12 996 women in the Avon Longitudinal Study of Parents and Children (ALSPAC) to develop prediction models and validated these in 3005 women from the Southampton Women's Survey (SWS). A model based on maternal early pregnancy characteristics only (BMI, height, age, parity, smoking, existing and previous gestational hypertension and diabetes, and ethnicity) plus initial mean arterial pressure was compared with a model additionally including current mean arterial pressure, a model including the deviation of current mean arterial pressure from a stratified normogram, and a model including both at different gestational ages from 20-36 weeks.

STUDY ANSWER AND LIMITATIONS
The addition of blood pressure measurements from 28 weeks onwards improved prediction models compared with use of early pregnancy risk factors alone, but they contributed little to the prediction of preterm birth or small for gestational age. Though multiple imputation of missing data was used to increase the sample size and minimise selection bias, the validation sample might have been slightly underpowered as the number of cases of pre-eclampsia was just below the recommended 100. Several risk factors were self reported, potentially introducing measurement error, but this reflects how information would be obtained in clinical practice.

WHAT THIS STUDY ADDS
The addition of routinely collected blood pressure measurements from 28 weeks onwards improves predictive models for pre-eclampsia based on blood pressure in early pregnancy and other characteristics, facilitating a reduction in scheduled antenatal care.

FUNDING, COMPETING INTERESTS, DATA SHARING
UK Wellcome Trust, US National Institutes of Health, and UK Medical Research Council. Other funding sources for authors are detailed in the full online paper. With the exceptions of CM-W, HMI, and KMG there were no competing interests.

Introduction
Pre-eclampsia is the most severe form of hypertensive disorder of pregnancy and is associated with maternal and fetal mortality, intrauterine growth restriction, and preterm birth. It is defined as de novo high blood pressure occurring after 20 weeks' gestation in conjunction with proteinuria, but these symptoms can develop swiftly, within days of a normal antenatal assessment, and deficiencies in care have previously been reported in relation to maternal and fetal deaths attributable to pre-eclampsia in the United Kingdom. Although there is no known cure for pre-eclampsia other than delivery of the baby, decisions such as whether to expectantly monitor or to deliver are important for both maternal and fetal outcome. Therefore it is important to be able to identify women at risk as early in pregnancy as possible to allow for adequate surveillance and early detection of disease. We, and others, have shown that before the diagnostic threshold for pre-eclampsia is met, the pattern of change in blood pressure during pregnancy differs between women who remain normotensive and those who subsequently develop pre-eclampsia. Specifically women who go on to experience pre-eclampsia have a higher initial blood pressure and a steeper increase after the mid-pregnancy nadir. A greater increase in blood pressure across pregnancy is also associated with a shorter gestation and a lower birthweight baby.

Several prediction models have shown the value of using blood pressure recorded at the first visit to
the antenatal clinic alongside other maternal characteristics to predict the later development of pre-eclampsia. As blood pressure is routinely measured on repeated occasions during pregnancy in most high income countries, and could also be feasibly measured in low and middle income countries, the incorporation of subsequent measurements and the trajectory of change in blood pressure in clinical prediction models could add additional value and improve the information available to make decisions about care. As well as identification of women at high risk who would benefit from intensive monitoring to prevent the adverse outcomes associated with pre-eclampsia, it might also be beneficial to be able to identify a group of women at low risk and transfer them to a reduced schedule of antenatal appointments.

We have previously developed stratified normal reference ranges for blood pressure, which allow for a different expected trajectory of blood pressure for different subgroups of women. In this study we assessed and externally validated the predictive ability of blood pressure measurements at different gestational ages, and the comparison with our reference ranges, to identify women at higher and low risk of pre-eclampsia using two prospective cohort studies from the UK. We additionally investigated whether such measurements were able to identify pregnancies at risk of preterm birth and small for gestational age babies, as these are two of the perinatal outcomes that are most commonly associated with pre-eclampsia.

This paper conforms to TRIPOD guidelines.

Methods

Avon Longitudinal Study of Parents and Children (ALSPAC)

For the development of the prediction models we used the Avon Longitudinal Study of Parents and Children (ALSPAC), a prospective birth cohort study that recruited women with expected delivery dates between 1 April 1991 and 31 December 1992 living in a defined area of Avon during pregnancy. The study has been described in full elsewhere, and the study website contains details of available data (www.bris.ac.uk/alspac/researchers/data-access/data-dictionary). A total of 14,541 women were enrolled, 13,678 had a singleton pregnancy resulting in a live birth, and data on measurements of blood pressure were available for 12996 of these women.

Obstetric measurements in ALSPAC

Six trained research midwives abstracted all measurements of blood pressure and proteinuria that were taken as part of routine antenatal care by midwives or obstetricians from the women’s obstetric records. There was no variation between midwives in mean values of the data abstracted, and error rates were consistently <1% in repeated data entry checks. Measurements were taken once in women in the seated position with the appropriate cuff size, with Korotkoff phase V used for diastolic blood pressure.

Maternal and offspring characteristics in ALSPAC

Maternal age at delivery and sex of the baby were obtained from obstetric records. Maternal weight and height before pregnancy, together with parity, smoking status, education, hypertension, and diabetes before the index pregnancy and ethnicity and household social class were obtained from questionnaires administered during pregnancy. Smoking during pregnancy was classified as “never” or “any smoking” for women who reported smoking either in the first three months of pregnancy or within two weeks of the 18 week questionnaire. Women who reported ever having hypertension outside of pregnancy were classed as having essential hypertension, while women who reported having hypertension only in pregnancy were classed as having previous gestational hypertension. Similarly, women who reported having ever had diabetes outside of pregnancy were defined as having diabetes while women who reported having diabetes only in a previous pregnancy were defined as having previous gestational diabetes.

Southampton Women’s Survey (SWS)

In accordance with recent guidelines on the development and assessment of prognostic models, we externally validated our prediction models in a further dataset. The Southampton Women’s Survey is a prospective cohort study that recruited non-pregnant women through general practices in Southampton between April 1998 and September 2002. Full details of the study have been published previously. The study enrolled 12,583 women, of whom 3158 became pregnant and had delivered a singleton liveborn infant by the end of 2007. Obstetric blood pressure measurements were available for 3005 of these.

Obstetric measurements in Southampton Women’s Survey

A team of trained research midwives abstracted all measurements of blood pressure that were routinely collected during pregnancy from the women’s obstetric records after delivery following a strict protocol. The measurements were ordered by time, and the data checked for consistency.

Maternal and offspring characteristics in Southampton Women’s Survey

Maternal age at delivery and sex of offspring were obtained from obstetric records. Information on maternal parity, education, social class, and ethnicity was obtained at the initial interview before pregnancy, and height was measured. At the 11 week interview the mothers reported pre-pregnancy weight, any diabetes in a previous pregnancy or outside of pregnancy, and high blood pressure in a previous pregnancy or outside of pregnancy. Smoking status was obtained during interviews at 11 and 34 weeks gestation and classified as “never smoked in pregnancy” or “any smoking in pregnancy.” An algorithm that incorporated information recorded before pregnancy and at six months postpartum was used to estimate smoking
in pregnancy when data at one or both time points were missing.

**Initial blood pressure and gestational age specific blood pressure in ALSPAC and SWS**

As preliminary analysis suggested that mean arterial pressure was a stronger predictor of each outcome than systolic blood pressure (SBP) or diastolic blood pressure (DBP) separately, we used this as our measure of blood pressure in prediction models. This was defined as \( (SBP + 2 \times DBP) / 3 \). In both cohorts, the initial measurement of mean arterial pressure was defined as the first antenatal measurement, provided that this occurred before 18 weeks’ gestation. We also derived mean arterial pressure at 20, 25, 28, 31, 34, and 36 weeks’ gestation for each woman (following the standard schedule of antenatal appointments for nulliparous women used in the UK)\(^6\), using the measurement that was closest to this gestational age and within two weeks of it, or classifying the variable as missing if there was no measurement within this gestational age range.

**Pregnancy outcomes in ALSPAC and SWS**

In ALSPAC pre-eclampsia was defined by the International Society for the Study of Hypertension in Pregnancy (ISSHP) criteria\(^6\) as systolic blood pressure \( \geq 140 \) mm Hg and/or diastolic blood pressure \( \geq 90 \) mm Hg along with proteinuria of \( \geq 1 \) + on urine dipstick testing occurring on two occasions after 20 weeks’ gestation. In SWS we used information on any diagnosis of pre-eclampsia that was recorded in obstetric records to define preeclampsia. In both cohorts, we defined small for gestational age as below the 10th centile of birth weight adjusted for gestational age at birth using internal standardisation by cohort. In ALSPAC the final clinical estimate of the expected date of delivery was abstracted from the obstetric records and used to calculate gestational age at delivery. In SWS we used a computerised algorithm to derive gestational age using menstrual data (66%) or, when these were uncertain or discrepant with ultrasound assessments, fetal anthropometry in early pregnancy. Preterm birth was defined as birth before 37 weeks’ gestation.

Table A in appendix 1 gives a summary of definitions of all outcomes and predictor variables in ALSPAC and SWS for comparison.

**Statistical analysis**

**Deriving expected trajectories of mean arterial pressure for each woman**

In the first stage of model development we developed expected trajectories of mean arterial pressure across pregnancy, or normograms, in ALSPAC, similarly to those described previously.\(^{24}\) For the 9,402 women in ALSPAC who had a normal pregnancy without pre-eclampsia, essential hypertension, or diabetes, and delivered a term infant appropriate size for gestational age we used multilevel models to describe the trajectory of change in mean arterial pressure using the observed repeated measurements across the whole of pregnancy and the exact gestational ages at measurement. The multilevel models had two levels: measurement occasion (level 1) within individual (level 2). We used restricted cubic splines with knots at 11, 18, 30, 36, and 40 weeks’ gestation to describe the shape of the trajectory as this was found to have the best fit to these data in previous analysis.\(^{26}\) We allowed for different average trajectories according to BMI category and smoking status by including appropriate interaction terms and fitted separate models for nulliparous and multiparous women. For all women in the ALSPAC cohort (including non-normal pregnancies) we then obtained a normative value of mean arterial pressure at 20, 25, 28, 31, 34, and 36 weeks’ gestation, which accounted for their parity and BMI before pregnancy and whether they smoked during pregnancy, and was conditional on their mean arterial pressure at their initial visit (before 18 weeks’ gestation). Appendix 2 gives further information on the derivation of this normative value.

**Missing data**

For both the ALSPAC and SWS datasets there were missing data on maternal characteristics and also for the derived values of initial mean arterial pressure and mean arterial pressure at 20, 25, 28, 31, 34, and 36 weeks for some women. To increase power and minimise selection bias in the second stage prediction models, we next used multivariable multiple imputation of missing data. Separate imputation models were used for the two cohorts, imputing to the full 12996 women in ALSPAC and 3005 women in SWS who had any measurements of mean arterial pressure in pregnancy. For each imputation model, we included all early pregnancy characteristics and derived mean arterial pressure values as well as the outcomes. We imputed any missing data for the normative values of mean arterial pressure passively from imputed early pregnancy characteristics and initial mean arterial pressure. For each imputation we generated 20 imputed datasets and combined coefficient estimates across these using Rubin’s rules.\(^{30, 31}\) Imputation of missing data in this way assumes that data are missing at random—that is, that any reasons for missingness are explained by the observed variables included in the imputation model.\(^{32}\) The assumption seems reasonable here as most women will not be aware of their current blood pressure (only their previous values) so this is unlikely to influence their decision to have their blood pressure measured at any given gestational age. Full details of the imputation models are included in appendix 2.

**Prediction of pre-eclampsia, preterm birth, and small for gestational age**

In the second stage we used logistic regression to obtain separate prediction models for pre-eclampsia, preterm birth, and small for gestational age in ALSPAC based on the imputed datasets. Maternal predictors were selected from BMI before pregnancy, height, age \( \geq 35 \), parity, smoking, essential hypertension, previous
gestational hypertension, diabetes, previous gestational diabetes, and non-white ethnicity. This set of predictors was chosen because of strong prior evidence of their association with pre-eclampsia, ease of measuring in clinical practice, and generalisability to different settings. We used backward selection, with a P value threshold of 0.2, to select variables for inclusion in these early pregnancy characteristic models. Mean arterial pressure at the initial visit was forced to be included in all of these prediction models. The early pregnancy characteristic model was called model 1. We then added mean arterial pressure at gestational ages 20, 25, 28, 31, 34, or 36 weeks (each tested separately) to determine the added value of such measurements from 20 weeks in predicting outcomes (model 2). In model 3 we added the difference between the woman’s measurement of mean arterial pressure at each gestational age and their normative value (that is, their deviation from their expected trajectory for their BMI, parity, smoking status, and initial mean arterial pressure) to model 1. Finally, model 4 combined models 2 and 3 by adding both the value of mean arterial pressure and the deviation from the normative value at each gestational age to model 1. Table B in appendix 1 gives a summary of variables included in each model. The areas under the receiver operating characteristics curves (ROCs) of the models were compared with a Wald test based on a bootstrapped standard error for the difference between areas. For predictions at each gestational age, we included in prediction models only women who had not yet delivered and aimed to predict the development of pre-eclampsia at any stage of pregnancy, any subsequent preterm births, and small for gestational age offspring measured at birth. Table C in appendix 1 shows the cumulative number of women who had delivered at each gestational age. We also considered including all measurements of mean arterial pressure up to a given gestational age in prediction models rather than just the initial and the current measure but found no meaningful improvement in the area under the curve.

The parameters from the first stage multilevel model were then used to derive normative values of mean arterial pressure at each gestational age in SWS and these were combined with parameters from the second stage logistic regression model developed in ALSPAC and applied to women in the SWS cohort to externally validate our prediction model. The model parameters for all outcomes are given in appendix 2.

We assessed the predictive ability of these models using the area under the receiver operating characteristic curve and the sensitivity, specificity, and positive and negative predictive values, together with positive and negative likelihood ratios. We assessed calibration by comparing the actual versus the predicted risk across 10ths of the risk score for each outcome. The prediction models were recalibrated for the SWS when necessary by adjusting the intercept and slope of the logistic regression model for this cohort (see appendix 2).33

We assessed the expected numbers of women out of an initial cohort of 1000 who would be classified as low risk and high risk if we were to use a sequential screening strategy to rule out pre-eclampsia at each gestational considered up to 31 weeks. Model 1 was used at booking, followed by model 2 at 20, 25, 28, and 31 weeks. We tested two thresholds for the screening strategy (applying the same threshold of the risk score at each gestational age of testing), which were derived from the average threshold (across all gestational ages from 20 to 36 weeks) required to give sensitivities of 95% and 99%, respectively, at each gestational age in ALSPAC. These thresholds were then applied to the SWS to produce a flowchart of the expected numbers of women in each group at each stage of testing.

All analyses were completed in Stata version 13.1 and multilevel models were fitted using MLwiN version 2.30 through Stata using the runmlwin command.34

Patient involvement

This paper is based on two prospective cohort studies that collect data from members of the general population (not patients) on a large number of potential risk factors and health outcomes and that are used widely by the research community to examine a wide range of health related questions. Consequently no patients were involved in setting the research question or the outcome measures, nor were they involved in recruitment or the design and implementation of the study. There are no plans to involve patients in dissemination.

Results

Table 1 shows maternal and pregnancy related characteristics in the imputed data pre-eclampsia status. Tables D and E in appendix 1 compare their distributions with observed data. Characteristics were generally similar in the imputed data compared with the observed data. Women in the SWS were more likely to be overweight or obese, to be nulliparous, to have had previous gestational hypertension or gestational diabetes, and to be of non-white ethnicity and were less likely to smoke or to have pre-existing hypertension, were shorter and older on average, and had higher initial mean arterial pressure than women in the ALSPAC cohort.

Table F in appendix 1 shows the unadjusted associations of each of the predictor variables with each of the outcomes.

Prediction of pre-eclampsia

Table 2 shows the area under the receiver operating characteristic curves for prediction of pre-eclampsia for the model with only early pregnancy characteristics and subsequent models adding mean arterial pressure at different gestational ages from 20 to 36 weeks. The prediction model that included only early pregnancy maternal characteristics and first mean arterial pressure from the antenatal clinic had an area under the curve of around 0.77 in the development cohort (ALSPAC; model 1). This improved to 0.79 with addition of mean arterial pressure (model 2) at...
Table 1 | Characteristics of pregnancies in Avon Longitudinal Study of Parents and Children (ALSPAC) and Southampton Women’s Survey (SWS) by pre-eclampsia status using imputed datasets

<table>
<thead>
<tr>
<th></th>
<th>ALSPAC Without pre-eclampsia (n=12,679)</th>
<th>With pre-eclampsia (n=317)</th>
<th>Overall (n=12,996)</th>
<th>SWS Without pre-eclampsia (n=2,917)</th>
<th>With pre-eclampsia (n=88)</th>
<th>Overall (n=3,005)</th>
</tr>
</thead>
<tbody>
<tr>
<td><strong>Mean (SD; range for overall mean)</strong></td>
<td></td>
<td></td>
<td></td>
<td><strong>Mean (SD; range for overall mean)</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Characteristics</strong></td>
<td></td>
<td></td>
<td></td>
<td><strong>Characteristics</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td><strong>Maternal BMI</strong></td>
<td></td>
<td></td>
<td></td>
<td><strong>Maternal BMI</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Underweight</td>
<td>5.9 (3.5)</td>
<td>3.5 (3.3)</td>
<td>5.9 (3.5)</td>
<td>2.7 (2.0)</td>
<td>3.1 (2.3)</td>
<td>2.7 (2.0)</td>
</tr>
<tr>
<td>Normal</td>
<td>72.7 (20.6)</td>
<td>60.6 (58.2)</td>
<td>72.4 (21.4)</td>
<td>63.8 (56.8)</td>
<td>44.0 (37.0)</td>
<td>57.6 (44.0)</td>
</tr>
<tr>
<td>Overweight</td>
<td>16.3 (15.6)</td>
<td>21.1 (19.5)</td>
<td>16.4 (15.8)</td>
<td>26.4 (23.8)</td>
<td>32.9 (26.1)</td>
<td>26.6 (23.1)</td>
</tr>
<tr>
<td>Obese</td>
<td>5.1 (4.6)</td>
<td>14.9 (14.0)</td>
<td>5.4 (4.7)</td>
<td>12.9 (11.6)</td>
<td>19.9 (18.0)</td>
<td>13.1 (11.3)</td>
</tr>
<tr>
<td>Mean (SD; range for overall mean) maternal height</td>
<td>163.9 (6.7)</td>
<td>163.0 (6.6)</td>
<td>163.9 (6.7)</td>
<td>163.9 (6.7)</td>
<td>164.3 (6.6)</td>
<td>163.2 (6.6)</td>
</tr>
<tr>
<td>Mean (SD; range for overall maternal age)</td>
<td>27.9 (5.0)</td>
<td>27.5 (4.9)</td>
<td>27.9 (5.0)</td>
<td>27.9 (5.0)</td>
<td>28.0 (4.9)</td>
<td>27.9 (4.9)</td>
</tr>
<tr>
<td>Maternal smoking in pregnancy (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>73.9 (73.0)</td>
<td>84.2 (84.3)</td>
<td>74.1 (74.3)</td>
<td>83.3 (83.3)</td>
<td>86.5 (86.5)</td>
<td>83.4 (83.4)</td>
</tr>
<tr>
<td>Yes</td>
<td>26.1 (27.0)</td>
<td>15.8 (15.7)</td>
<td>25.8 (26.0)</td>
<td>16.7 (16.7)</td>
<td>13.5 (13.5)</td>
<td>16.6 (16.6)</td>
</tr>
<tr>
<td>Essential hypertension (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>96.5 (96.5)</td>
<td>88.4 (88.4)</td>
<td>96.3 (96.3)</td>
<td>97.6 (97.6)</td>
<td>88.5 (88.5)</td>
<td>97.4 (97.4)</td>
</tr>
<tr>
<td>Yes</td>
<td>3.5 (3.5)</td>
<td>11.6 (11.6)</td>
<td>3.7 (3.7)</td>
<td>2.4 (2.4)</td>
<td>11.5 (11.5)</td>
<td>2.6 (2.6)</td>
</tr>
<tr>
<td>Previous gestational hypertension (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>91.2 (91.2)</td>
<td>84.7 (84.7)</td>
<td>91.0 (91.0)</td>
<td>63.7 (63.7)</td>
<td>18.4 (18.4)</td>
<td>62.4 (62.4)</td>
</tr>
<tr>
<td>Yes</td>
<td>8.8 (8.8)</td>
<td>15.3 (15.3)</td>
<td>9.0 (9.0)</td>
<td>36.3 (36.3)</td>
<td>81.6 (81.6)</td>
<td>37.6 (37.6)</td>
</tr>
<tr>
<td>Diabetes (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>99.7 (99.7)</td>
<td>96.1 (96.1)</td>
<td>99.6 (99.6)</td>
<td>99.5 (99.5)</td>
<td>96.8 (96.8)</td>
<td>99.4 (99.4)</td>
</tr>
<tr>
<td>Yes</td>
<td>0.3 (0.3)</td>
<td>3.9 (3.9)</td>
<td>0.4 (0.4)</td>
<td>0.5 (0.5)</td>
<td>3.2 (3.2)</td>
<td>0.6 (0.6)</td>
</tr>
<tr>
<td>Previous gestational diabetes (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No</td>
<td>99.5 (99.5)</td>
<td>98.8 (98.8)</td>
<td>99.5 (99.5)</td>
<td>92.0 (92.0)</td>
<td>78.8 (78.8)</td>
<td>91.7 (91.7)</td>
</tr>
<tr>
<td>Yes</td>
<td>0.5 (0.5)</td>
<td>1.2 (1.2)</td>
<td>0.5 (0.6)</td>
<td>8.0 (8.0)</td>
<td>21.2 (21.2)</td>
<td>8.3 (8.3)</td>
</tr>
<tr>
<td>Maternal ethnicity (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>White</td>
<td>97.3 (97.3)</td>
<td>95.8 (95.8)</td>
<td>97.2 (97.2)</td>
<td>95.4 (95.4)</td>
<td>97.7 (97.7)</td>
<td>95.4 (95.4)</td>
</tr>
<tr>
<td>Non-white</td>
<td>2.7 (2.7)</td>
<td>4.2 (4.2)</td>
<td>2.8 (2.8)</td>
<td>4.6 (4.6)</td>
<td>2.3 (2.3)</td>
<td>4.6 (4.6)</td>
</tr>
</tbody>
</table>
20 weeks of gestation and became greater with increasing gestational age at which mean arterial pressure was measured, being 0.88 at 36 weeks of gestation. In the validation cohort (SWS) areas under the curve were similar to those in the development cohort, but the earliest gestational age at which there was evidence of an improvement in the area under the curve for model 2 compared with model 1 was at 28 weeks. We therefore focused on this gestational age when presenting calibration and predictive statistics. Models 3 and 4 performed similarly to model 2 at all gestational ages in ALSPAC and SWS (table G in appendix 1). Figure A in appendix 3 shows the ROC curves for all four models with mean arterial pressure measured at 28 weeks in ALSPAC and SWS.

Table 2 | Prediction models for pre-eclampsia at different gestational ages in Avon Longitudinal Study of Parents and Children (ALSPAC) and Southampton Women’s Survey (SWS) using imputed datasets. Figures are area under receiver operating characteristic curve (95% confidence intervals)

<table>
<thead>
<tr>
<th>Gestational age of prediction (weeks)</th>
<th>Development cohort (ALSPAC)</th>
<th>Validation cohort (SWS)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Model 1*</td>
<td>Model 2†</td>
</tr>
<tr>
<td>20</td>
<td>0.77 (0.75 to 0.80)</td>
<td>0.79 (0.77 to 0.82)</td>
</tr>
<tr>
<td>25</td>
<td>0.77 (0.75 to 0.80)</td>
<td>0.80 (0.77 to 0.82)</td>
</tr>
<tr>
<td>28</td>
<td>0.77 (0.75 to 0.80)</td>
<td>0.81 (0.79 to 0.84)</td>
</tr>
<tr>
<td>31</td>
<td>0.77 (0.74 to 0.79)</td>
<td>0.82 (0.80 to 0.85)</td>
</tr>
<tr>
<td>34</td>
<td>0.77 (0.74 to 0.80)</td>
<td>0.84 (0.82 to 0.87)</td>
</tr>
<tr>
<td>36</td>
<td>0.77 (0.74 to 0.80)</td>
<td>0.88 (0.86 to 0.90)</td>
</tr>
</tbody>
</table>

*Includes maternal covariates and mean arterial pressure at first visit.
†As model 1 plus observed value of mean arterial pressure for gestational age.

Table H in appendix 1 and figure 1 show the calibration of model 2 at 28 weeks’ gestation for prediction of pre-eclampsia at 10ths of the risk score in the validation cohort (SWS). The model tended to overestimate the risk of pre-eclampsia, especially at higher values of the risk score, suggesting that we needed to recalibrate the model. Recalibration details are given in the appendix 2. After recalibration the risk of pre-eclampsia was well estimated by the model in the validation cohort (table H in appendix 1; fig 1).

Table 3 shows specificity, positive and negative predictive values, and positive and negative likelihood ratios at fixed sensitivities of 95% and 99% (meaning that 95% and 99% of women who developed pre-eclampsia would test positive) in ALSPAC and SWS for model 2 at 28 weeks. At these high sensitivities, as expected, there was a high false positive rate (low specificity), though we also observed a high negative predictive value (about 1) and low negative likelihood ratio, suggesting a strong ability to rule out pre-eclampsia in women who tested negative at these values of the risk score. Figure 2 shows the expected results of a sequential screening strategy to rule out pre-eclampsia, testing with model 1 at booking and then with model 2 from 20 weeks onwards, with the threshold of the risk score set according to a sensitivity of 99% at each gestational age in ALSPAC, and the threshold applied in SWS. Women are moved into the low risk group at any gestational age at which they test negative for pre-eclampsia, and once in the low risk group women are not included in subsequent screens. Based on this approach it is expected that at 31 weeks, 27 out of 28 women who had not yet delivered and would go on to develop pre-eclampsia would be included in the high risk group and that just over a third of all of the women would have been allocated to the low risk group with a reduced care schedule. Figure B in appendix 3 shows the equivalent result with the threshold of the risk score that gave a 95% sensitivity in ALSPAC.

Table I in appendix 1 shows predictive statistics fixing low false positive rates of 25%, 10%, and 5% (high specificities) for model 2 at 28 weeks. For a 5% false positive rate the sensitivity was 35% in ALSPAC and 32% in SWS, suggesting that the model had a weaker ability to rule in pre-eclampsia than to rule it out.
**Prediction of pre-eclampsia in Avon Longitudinal Study of Parents and Children (ALSPAC) and Southampton Women’s Survey (SWS) with model 2* at 28 weeks’ gestation for fixed sensitivities using imputed datasets**

<table>
<thead>
<tr>
<th>Sensitivity</th>
<th>Risk score threshold</th>
<th>Specificity (95% CI)</th>
<th>Positive predictive value (95% CI)</th>
<th>Negative predictive value (95% CI)</th>
<th>Positive likelihood ratio (95% CI)</th>
<th>Negative likelihood ratio (95% CI)</th>
<th>No (%) testing positive</th>
<th>No (%) true positive</th>
<th>No (%) false positive</th>
<th>No (%) true negative</th>
<th>No (%) false negative</th>
</tr>
</thead>
<tbody>
<tr>
<td>ALSPAC</td>
<td>0.95</td>
<td>0.008</td>
<td>0.37 (0.31 to 0.42)</td>
<td>1.00 (0.99 to 1.00)</td>
<td>1.50 (1.37 to 1.64)</td>
<td>0.14 (0.08 to 0.23)</td>
<td>8316 (64.1)</td>
<td>8019 (61.9)</td>
<td>4673 (35.7)</td>
<td>16 (0.1)</td>
<td>297 (2.3)</td>
</tr>
<tr>
<td>SWS (without recalibration, using ALSPAC risk score thresholds)</td>
<td>0.99</td>
<td>0.003</td>
<td>0.15 (0.13 to 0.17)</td>
<td>1.00 (1.00 to 1.00)</td>
<td>1.16 (1.13 to 1.20)</td>
<td>0.07 (0.02 to 0.21)</td>
<td>11068 (85.4)</td>
<td>310 (2.4)</td>
<td>10758 (83.0)</td>
<td>1894 (14.6)</td>
<td>3 (0.0)</td>
</tr>
<tr>
<td>SWS (after recalibration)</td>
<td>0.96</td>
<td>0.011</td>
<td>0.47 (0.40 to 0.55)</td>
<td>1.00 (0.99 to 1.00)</td>
<td>1.44 (1.40 to 1.50)</td>
<td>0.02 (0.00 to 0.32)</td>
<td>2086 (69.6)</td>
<td>87 (2.9)</td>
<td>2000 (66.7)</td>
<td>913 (30.4)</td>
<td>0 (0.0)</td>
</tr>
</tbody>
</table>

*S includes maternal covariates, mean arterial pressure at first visit, and observed value of mean arterial pressure at 28 weeks.

†Continuity correction applied in calculation of likelihood ratios in multiply imputed datasets where 2×2 contingency tables of test vs disease status had zero cells. Constant of 0.5 has been added to all cells in these cases.
Study strengths and limitations

We used two large prospective cohort studies, one to develop the prediction models and the second to externally validate them. The detailed measurements of blood pressure in each of the studies allowed us to evaluate the predictive ability of these measurements at multiple gestational ages across pregnancy. We also used multiple imputation to maximise the sample size available. The ALSPAC pregnancies occurred over 20 years ago, and, as then, there might have been changes in clinical practice. The models, however, were validated in the SWS, where the pregnancies occurred more recently (8-16 years ago). Several of the predictor variables used were self reported so could be prone to measurement error. This is likely to reflect the way that information is obtained in clinical practice. We needed to recalibrate the prediction model for pre-eclampsia in the SWS. This is probably because of different definitions of pre-eclampsia in the two cohorts; in ALSPAC pre-eclampsia was defined by applying a standard definition to all measurements of blood pressure and proteinuria in pregnancy whereas in SWS a record of a clinical diagnosis of pre-eclampsia during pregnancy was used. This suggests that the risk score requires recalibration according to different definitions of pre-eclampsia. The areas under the curve, however, were similar in the SWS to those in ALSPAC suggesting that the predictive ability of measurements of blood pressure in the second half of pregnancy is not dependent on the use of a particular definition of pre-eclampsia. The number of cases of pre-eclampsia in the SWS (n=88) was slightly lower than the recommended 100 events for validation,35 and our study might therefore have been slightly underpowered to detect differences in model fit. Despite this, our confidence intervals were of a reasonable width, and we were able to detect both the need for recalibration in SWS and an improvement in the area under the curve for pre-eclampsia from the addition of blood pressure measures from 28 weeks onwards in this cohort. We could not distinguish between medically indicated and spontaneous preterm births in our analyses, but our sensitivity analysis excluding women who had pre-eclampsia is likely to exclude most medically indicated preterm births related to high blood pressure.

Comparisons with other studies

The area under the curve for our prediction model for pre-eclampsia using mean arterial pressure at 28 weeks was 0.81 in ALSPAC and 0.83 in SWS. This is higher than the area found in a model based on maternal clinical risk factors (including mean arterial pressure) in nulliparous women only in the SCOPE study at 15 weeks’ gestation (0.76 in training dataset and 0.71 after validation).

Table M in appendix 1 shows the risk factors included in that model (not all of these were available in ALSPAC and SWS for us to compare our model directly with the SCOPE model). A recent systematic review identified 69 prediction models for pre-eclampsia but reported that only five of these had been externally validated.36 The validated models all included uterine artery Doppler measures or biomarkers that require laboratory testing that is not currently part of routine practice and is too expensive for use in low or middle income countries.21-24 Our prediction model uses readily available measures and allows for a sequential screening approach at each antenatal appointment that could be used in conjunction with screening in early pregnancy to update each woman’s risk estimate as pregnancy progresses. In a meta-analysis, mean arterial pressure in the second trimester was found to predict pre-eclampsia well, with an area under the curve of 0.76 in low risk populations.41 We have shown that this increases with gestation and that mean arterial pressure in the second half of pregnancy contributes additional predictive ability compared with using early pregnancy clinical characteristics alone.

Conclusions and policy implications

Current models of antenatal care focus on a series of repeat scheduled antenatal visits with recurrent screening for pre-eclampsia. In the UK, nulliparous
women will have 12 healthcare assessments across gestation with nine of these visits occurring after 20 weeks’ gestation, while multiparous women will have nine assessments with six occurring after 20 weeks, missing appointments at 25, 31, and 40 weeks.\(^2\) Similar stratified (based on parity) care occurs in other Western countries. A potential use of our model is for nulliparous women who are screened as low risk to be transferred to a reduced schedule similar to that of a multiparous woman. Conversely, it might be appropriate to transfer multiparous women screening as high risk to the more intensive schedule of a nulliparous woman. We have provided two examples of how care schedules could be altered in this way following sequential screening (fig 3). The decision of whether to increase the schedule of antenatal appointments for women testing as high risk versus decreasing the schedule for women classified as low risk, however, will depend on the setting and resources available for screening. A systematic review of randomised controlled trials that compared standard antenatal care with reduced care schedules found no difference in the rates of detection of pre-eclampsia in studies of low risk women in high income countries.\(^3\) Nevertheless, in low income countries the standard schedule of antenatal care might consist of only four visits across pregnancy, and a low proportion of women achieve full attendance at these; here a focus on increasing the number of visits or increasing the awareness of the importance of these visits for women who screen as high risk might be most beneficial.\(^4\) It is also important to note that we have not taken into account reasons for monitoring pregnant women other than pre-eclampsia, and these must also be considered when this care schedule is planned.

Further qualitative research is required to evaluate the views of clinicians and pregnant women on the frequency of antenatal care as many women find this regular contact with healthcare professionals reassuring during pregnancy.\(^5\) This would inform the potential benefits and harms of different screening strategies and could be followed up with decision curve analysis to weigh these against each other.\(^6\) If the prediction models were to be used widely this would require the development of a tool for clinical use. Such a tool could be incorporated into a smart phone app, and future research could focus on evaluating the use of this tool in low and middle income settings, in a similar way to the development and evaluation of tools for smartphones that has occurred in ophthalmology.\(^7,\)\(^8\) Finally, extension of the models with additional measures such as the number and gestation of previous preterm deliveries or cervical length for preterm birth\(^9\) or first trimester serum concentrations of pregnancy proteins for small for gestational age,\(^10,\)\(^11\) given that placental dysfunction is likely to underlie growth restriction, might be able to improve prediction of the adverse outcomes associated with pre-eclampsia.

In conclusion, measurements of blood pressure recorded during the second half of pregnancy, used in conjunction with blood pressure early in pregnancy and other maternal risk factors, can improve the identification of women who are at risk of developing pre-eclampsia later in pregnancy and could be used to differentiate women who require more intensive monitoring from those who are likely to have a normal pregnancy. As there is little value in additionally using these later pregnancy measurements to predict preterm birth and small for gestational age, however, disease specific models will be required to facilitate stratified care for the range of pregnancy complications.

**AUTHOR AFFILIATIONS**

\(^1\)MRC Integrative Epidemiology Unit at the University of Bristol, Bristol BS8 2BN, UK
\(^2\)School of Social and Community Medicine, University of Bristol, Bristol BS8 2BN, UK
\(^3\)Department of Medical Statistics, London School of Hygiene and Tropical Medicine, London WC1E 7HT, UK
\(^4\)Centre for Statistical Methodology, London School of Hygiene and Tropical Medicine, London WC1E 7HT, UK

---

**Example 1:** Nulliparous woman who screens below threshold of risk prediction score for pre-eclampsia from 20 weeks onwards
Usual schedule of appointments: Booking, 16 weeks, 20 weeks, 28 weeks, 31 weeks, 34 weeks, etc

<table>
<thead>
<tr>
<th>Screened (model 1)</th>
<th>Screened (model 1)</th>
<th>Screened (model 2)</th>
<th>Not screened</th>
<th>Screened (model 2)</th>
<th>Not screened</th>
<th>Screened (model 2)</th>
<th>Not screened</th>
</tr>
</thead>
<tbody>
<tr>
<td>Booking</td>
<td>Above threshold - continue with usual schedule</td>
<td>Above threshold - continue with usual schedule</td>
<td>Below threshold - transfer to reduced schedule and skip next appointment</td>
<td>20 weeks</td>
<td>Below threshold - transfer to reduced schedule and skip next appointment</td>
<td>25 weeks</td>
<td></td>
</tr>
<tr>
<td></td>
<td>16 weeks</td>
<td></td>
<td>28 weeks</td>
<td></td>
<td>31 weeks</td>
<td></td>
<td>34 weeks</td>
</tr>
<tr>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

**Example 2:** Multiparous woman who screens above threshold of risk prediction score for pre-eclampsia from 20 weeks gestation, but below at 28 weeks gestation
Usual schedule of appointments: Booking, 16 weeks, 20 weeks, 28 weeks, 34 weeks, etc

<table>
<thead>
<tr>
<th>Screened (model 1)</th>
<th>Screened (model 1)</th>
<th>Screened (model 2)</th>
<th>Screened (model 2)</th>
<th>Not screened</th>
<th>Screened (model 2)</th>
<th>Not screened</th>
<th>Screened (model 2)</th>
<th>Not screened</th>
</tr>
</thead>
<tbody>
<tr>
<td>Booking</td>
<td>Above threshold - continue with usual schedule</td>
<td>Above threshold - continue with usual schedule</td>
<td>Above threshold - transfer to more intensive schedule and add in extra appointment at 28 weeks</td>
<td>20 weeks</td>
<td>Above threshold - remain on more intensive schedule</td>
<td>25 weeks</td>
<td>Below threshold - transfer back to reduced schedule with no appointment at 31 weeks</td>
<td>28 weeks</td>
</tr>
<tr>
<td></td>
<td>16 weeks</td>
<td></td>
<td></td>
<td>25 weeks</td>
<td></td>
<td>28 weeks</td>
<td></td>
<td>31 weeks</td>
</tr>
</tbody>
</table>
RESEARCH

1MRC Lifecourse Epidemiology Unit, University of Southampton, Southampton General Hospital, Southampton SO16 6YD, UK
2NIHR Southampton Biomedical Research Centre, University of Southampton and University Hospital Southampton NHS Foundation Trust, Southampton General Hospital, Southampton SO16 6YD, UK
3National Institute for Health Research Musculoskeletal Biomedical Research Unit, University of Oxford, Oxford OX3 4LE, UK